請依序作答。

一、單選題, 回答正確選項即可, 不需說明。 每題3分, 共45分。

- 1. In the linear probability model, the interpretation of the slope coefficient is
	- (a) the change in odds associated with a unit change in X , holding other regressors constant.
	- (b) not all that meaningful since the dependent variable is either o or 1.
	- (c) the change in probability that $Y = 1$ associated with a unit change in X, holding others regressors constant.
	- (d) the response in the dependent variable to a percentage change in the regressor.
- 2. The maximum likelihood estimation method produces, in general, all of the following desirable properties with the **exception** of
	- (a) efficiency.
	- (b) consistency.
	- (c) normally distributed estimators in large samples.
	- (d) unbiasedness in small samples.
- 3. Having more relevant instruments
	- (a) is a problem because instead of being just identified, the regression now becomes overidentified.
	- (b) is like having a larger sample size in that the more information is available for use in the IV regressions.
	- (c) typically results in larger standard errors for the TSLS estimator.
	- (d) is not as important for inference as having the same number of endogenous variables as instruments.
- 4. Weak instruments are a problem because
	- (a) the TSLS estimator may not be normally distributed, even in large samples.
	- (b) they result in the instruments not being exogenous.
	- (c) the TSLS estimator cannot be computed.
	- (d) you cannot predict the endogenous variables any longer in the first stage.
- 5. Failure to follow the treatment protocol means that
	- (a) the OLS estimator cannot be computed.
	- (b) instrumental variables estimation of the treatment effect should be used where the initial random assignment is the instrument for the treatment actually received.
	- (c) you should use theTSLS estimator and regress the outcome variableYon the initialrandom assignment in the first stage to get predicted values of the outcome variable.
	- (d) the Hawthorne effect plays a crucial role.
- 6. In a sharp regression discontinuity design:
	- (a) crossing the threshold influences receipt of the treatment but is not the sole determinant.
	- (b) the population regression line must be linear above and below the threshold.
	- (c) X_i will in general be correlated with u_i .
	- (d) receipt of treatment is entirely determined by whether W exceeds the threshold.
- 7. In the Project STAR experiment, the observations are not plausibly i.i.d. As a result,
	- (a) we should use homoskedasticity-only standard errors.
	- (b) we should use clustered standard errors at the school level.
	- (c) we should use clustered standard errors across schools.
	- (d) the OLS estimator will be biased.
- 8. The MSPE by m -fold cross validation is given by the following formula:
	- (a) $\widehat{MSPE}_{m-fold\ cross\ validation} = \frac{1}{m} \sum_{i=1}^{m} \widehat{MSPE}_{i}$.
	- (b) $\widehat{MSPE}_{m-fold\ cross\ validation} = \sum_{i=1}^{m} \left(\frac{n_i}{n_i} \right)$ $\left(\frac{n_i}{n/m}\right)$ \widehat{MSPE}_i .
	- (c) $\widehat{MSPE}_{m-fold\ cross\ validation} = \frac{1}{m} \sum_{i=1}^{m} \left(\frac{n_i}{n/i} \right)$ $\left(\frac{n_i}{n/m}\right)^2 \widehat{MSPE}_i.$
	- (d) $\widehat{MSPE}_{m-fold\ cross\ validation} = \frac{1}{m} \sum_{i=1}^{m} \left(\frac{n_i}{n_i} \right)$ $\left(\frac{n_i}{n/m}\right)$ \widehat{MSPE}_i .
- 9. In order for the ridge regression estimator to work well for the data at hand, you choose λ_{Ridge} (a) arbitrarily.
	- (b) by minimizing $S^{Ridge}(b, \lambda_{Ridge})$.
	- (c) by setting it to 1.
	- (d) to minimize the estimated MSPE.
- 10. When k = 1, the Lasso estimator is given by the following formula when $\hat{\beta} \geq$ 0:

(a)
$$
\hat{\beta}^{Lasso} = \min \left(\hat{\beta} - \frac{1}{2} \frac{\lambda^{Lasso}}{\sum_{i=1}^{n} X_i^2}, 0 \right)
$$

(b)
$$
\hat{\beta}^{Lasso} = \max \left(\hat{\beta} - \frac{1}{2} \frac{\lambda^{Lasso}}{\sum_{i=1}^{n} X_i^2}, o \right)
$$

(c)
$$
\hat{\beta}^{Lasso} = \min \left(\hat{\beta} + \frac{1}{2} \frac{\lambda^{Lasso}}{\sum_{i=1}^{n} X_i^2}, 0 \right)
$$

(d)
$$
\hat{\beta}^{Lasso} = \max \left(\hat{\beta} + \frac{1}{2} \frac{\lambda^{Lasso}}{\sum_{i=1}^{n} X_i^2}, o \right)
$$

- 11. The BIC is a statistic
	- (a) used to help the researcher choose the number of lags in an autoregression.
	- (b) commonly used to test for serial correlation.
	- (c) only used in cross-sectional analysis.
	- (d) developed by the Bank of England in its river of blood analysis.
- 12. You should use the QLR test for breaks in the regression coefficients, when
	- (a) the Chow F -test has a p value of between 0.05 and 0.10.
	- (b) the suspected break data is not known.
	- (c) there are breaks in only some, but not all, of the regression coefficients.
	- (d) the suspected break data is known.
- 13. Problems caused by stochastic trends include all of the following with the **exception** of
	- (a) the estimator of an $AR(1)$ is biased towards zero if its true value is one.
	- (b) the model can no longer be estimated by OLS.
	- (c) t -statistics on regression coefficients can have a non-normal distribution, even in large samples.
	- (d) the presence of spurious regression.
- 14. HAC standard errors should be used because
	- (a) they are convenient simpliûcations of the heteroskedasticity-robust standard errors.
	- (b) conventional standard errors may result in misleading inference.
	- (c) they are easier to calculate than the heteroskedasticity-robust standard errors and yet still allow you to perform inference correctly.
	- (d) when there is a structural break, then conventional standard errors result in misleading
- 15. Consider the distributed lag model $Y_t = \beta_0 + \beta_1 X_t + \beta_2 X_{t-1} + \beta_3 X_{t-2} + \cdots + \beta_{r+1} X_{t-r} + u_t$. The dynamic causal effect is
	- (a) $\beta_0 + \beta_1$.
	- (b) $\beta_1 + \beta_2 + \cdots + \beta_{r+1}$.
	- (c) $\beta_0 + \beta_1 + \cdots + \beta_{r+1}$.
	- (d) β_1 .

二、非選擇題, 共55分。 答題時請務必簡要清楚, 並適當說明你的想法, 答案以能讓閱卷者瞭解為原則。

1. $(20%)$ Consider a model where random variables X and Y are jointly determined by the following equations:

$$
Y = \beta_1 + \beta_2 X + \beta_3 Z + u,\tag{1}
$$

$$
X = \gamma_1 + \gamma_2 Y + \nu, \tag{2}
$$

where *u* and *v* are error terms, mutually uncorrelated with mean zero and variance σ_u^2 and σ_v^2 , and Z is an exogenous variable.

- (a) (2%) Are β_2 and γ_2 identified? Explain.
- (b) (2%) Derive the covariance between Y and v, Cov(Y, v).
- (c) (4%) Is the OLS estimate of y_2 in equation (2) consistent in general? Is it consistent if $\beta_2 = 0$?
- (d) (4%) Is there any valid instrument to estimate y_2 . Why? If yes, show that the corresponding instrument variable estimate γ_2^{IV} $i₂^{IV}$ is consistent.
- (e) (4%) Show that the two stage least square (2SLS) estimate of y_2 is identical to the IV estimate in (d). Note that that 2SLS estimate is the OLS estimate of γ_2 in (2) when Y is replace by its fitted values by OLS in the equation $Y = \delta_1 + \delta_2 Z + w$, where w is an error term.
- (f) (4%) Suppose that $y_1 =$ o. Is $\frac{\bar{X}}{\bar{Y}}$ $\frac{\tilde{X}}{\tilde{Y}}$ a consistent estimator of γ_2 ? Why?
- 2. (10%) Suppose Y_t follows a random walk, $Y_t = Y_{t-1} + u_t$, for $t = 1, \dots, T$, where $Y_0 = 0$ and u_i is i.i.d. with mean 0 and variance σ_u^2 .
	- (a) (4%) Compute the mean and variance of Y_t .
	- (b) (3%) Compute the covariance of Y_t and Y_{t-k} .
	- (c) (3%) Use the results in (a) and (b) to determine whether Y_t is stationary or nonstationary.

3. (10%) Suppose there are panel data for $T = 2$ time periods for a randomized controlled experiment, where the first observation ($T = 1$) is taken before the experiment and the second observation $(t = 2)$ is for the posttreatment period. Suppose the treatment is binary; that is, suppose $X_{it} = 1$ if the *i*th individual is in the treatment group and $t = 2$, and $X_{it} = 0$ otherwise. Further suppose the treatment effect can be modeled using the specification

$$
Y_{it} = \alpha_i + \beta_1 X_{it} + u_{it},
$$

where α_i are individual-specific effects with a mean of 0 and a variance of σ_{α}^2 and u_{it} is an error term, where u_{it} is **homoskedastic**, $Cov(u_{i1}, u_{i2}) = o$, and $Cov(u_{it}, \alpha_i) = o$ for all *i*. Let $\hat{\beta}_1^{differences}$ 1 denote the differences estimator— that is, the OLS estimator in a regression of Y_{i2} on X_{i2} with an intercept— and let $\hat{\beta}_1^{diffs-in-diffs}$ denote the differences-in-differences estimator— that is, the estimator of β_1 based on the OLS regression of $\Delta Y_i = Y_{i2} - Y_{i1}$ against $\Delta X_i = X_{i2} - X_{i1}$ and an intercept. (**Note**: The homoskedasticity-only formula for the variance of the OLS estimator of β_1 for $Y_i = \beta_o + \beta_1 X_i + \epsilon$ is $\text{Var}(\hat{\beta}_1) = \frac{\sigma_{\epsilon}^2}{n \text{Var}(X_i)}$.)

- (a) (4%) What is the variance of the differences estimator, $Var(\hat{\beta}_1^{differences})$?
- (b) (4%) What is the variance of the differences-in-differences estimator, Var($\hat{\beta}_1^{diffs-in-diffs}$)?
- (c) $(2%)$ Based on your answers to (a) and (b) , when would you prefer the differences-indifferences estimator over the differences estimator, based purely on efficiency consideration.
- 4. (15%) A prospective employer receives two resumes: a resume from a **white** job applicant and similar resume from an **African American** applicant. Is the employer more likely to call back the white applicant to arrange an interview? Bertrand and Mullainathan (2004) carried out a randomized controlled experiment to answer this question. Because race is not typically included on a resume, they differentiated resumes on the basis of "white-sounding names" and "African American-sounding names". A large collection of fictitious resumes was created, and the presupposed "race" (base on the "sound" of the name) was randomly assigned to each resume. These resumes were sent to prospective employers to see which resumes generated a phone call (a **call back**) from prospective employer. Refer to the attached pages for "**Document for Names Data**" and various regression results from the **STATA log ûle**, answer the following questions.
	- (a) $(4%)$ Define the *callback rate* as the fraction of resumes that generate a phone call from the prospective employer. What was the call back rate for whites? For African Americans? Is the difference statistically significant? Is it large in a real-world sense?
	- (b) $(3%)$ Is the African American/white callback rate differential different for men than for women?
	- (c) (4%) What was the callback rate for high-quality resumes? For low-quality resumes? Is there a significant difference in callback rates for **high-quality** versus **low-quality** resumes?
	- (d) $(4%)$ What is the high-quality/low-quality difference for white applicants? For African American applicants? Is there a significant difference in this high-quality/low-quality differences for whites versus African Americans?

Documentation for Names Data

Names contains resume, call-back and employer information for 4,870 fictitious resumes sent in response to employment advertisements in Chicago and Boston in 2001, in a randomized controlled experiment conducted by Marianne Bertrand and Sendhil Mullainathan. The resumes contained information concerning the race of the applicant. Because race is not typically included on a resume, resumes were differentiated on the basis of so-called "white sounding names" (such as Emily Walsh or Gregory Baker) and "African American sounding names" (such as Lakisha Washington or Jamal Jones). A large collection of fictitious resumes were created and the presupposed "race" (based on the "sound" of the name) was randomly assigned to each resume. These resumes were sent to prospective employers to see which resumes generated a phone call (a "call back") from the prospective employer. These data were provided by Professor Marianne Bertrand of the University of Chicago, and were used in her paper with Sendhil Mullainathan "Are Emily and Greg More Employable that Lakisha and Jamal? A Field Experiment on Labor Market Discrimination," *American Economic Review*" 2004, Vol. 94, no. 4.

Variable Descriptions are provided on the next page.

Variable Descriptions

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 name: <unnamed> log: final_names.log

log type: text

opened on: 17 Jun 2021, 11:34:01

. use names.dta

. desc

.

Contains data from names.dta obs: 4,870

Sorted by:

. summarize

.

. $/*$ model (4) $*/$. reg call back black high h b, r

