# Dealing with Heterogeneity: Finite Mixture Models 處理群體異質性: 有限混入模型

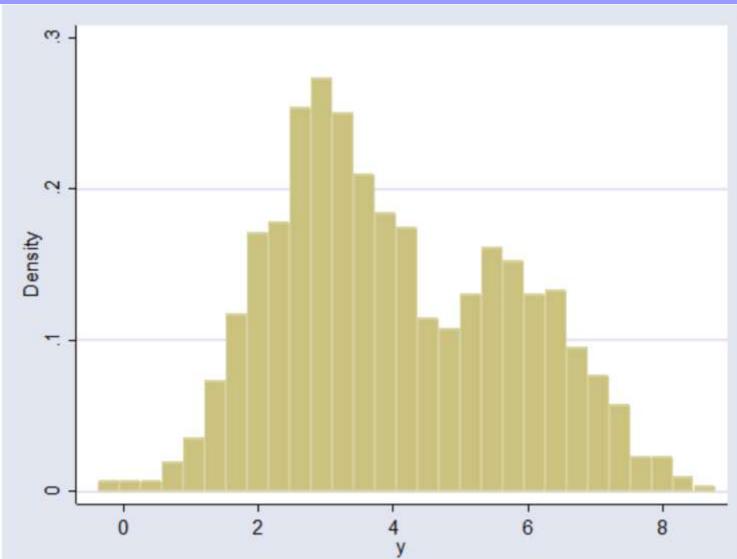
#### Joseph Tao-yi Wang (王道一) Experimetrics Lecture 6 (實驗計量第六講)

# Part I: Mixture of Two Normal Distributions 第一部分: 混入兩個常態分配

#### Joseph Tao-yi Wang (王道一) Experimetrics Lecture 6 (實驗計量第六講)

### Mixture of Two Normal Distributions

- Data (N=1,000)
  mixture\_sim.dta
- STATA Command: hist y
- STATA Results:
  2 Types of Subjects?
  Mean at 3 and 6?



## Mixture of Two Normal Distributions

# Type 1:

- Mixing Proportion Pr(Type 1) = p
- Choose  $y \sim N(\mu_1, \sigma_1^2)$  with  $f(y|\text{Type 1}) = \frac{1}{\sigma_1}\phi\left(\frac{y-\mu_1}{\sigma_1}\right)$

# Type 2:

- Mixing Proportion Pr(Type 2) = (1 p)Choose  $y \sim N(\mu_2, \sigma_2^2)$  with  $f(y|\text{Type 2}) = \frac{1}{\sigma_2}\phi\left(\frac{y \mu_2}{\sigma_2}\right)$
- Marginal Density (Likelihood):

$$f(y;\underline{\mu_1,\sigma_1,\mu_2,\sigma_2,p}) = p \cdot \frac{1}{\sigma_1} \phi\left(\frac{y-\mu_1}{\sigma_1}\right) + (1-p) \cdot \frac{1}{\sigma_2} \phi\left(\frac{y-\mu_2}{\sigma_2}\right)$$

### Mixture of Two Normal Distributions

• Estimate  $\hat{\mu_1}, \hat{\sigma_1}, \hat{\mu_2}, \hat{\sigma_2}, \hat{p}$  to max. n

Sample log-Likelihood:  $\log L = \sum_{i=1} \ln f(y_i; \mu_1, \sigma_1, \mu_2, \sigma_2, p)$ (for  $y_1, y_2, \dots, y_n$ )

Calculate Posterior Probability:

$$\Pr(\text{Type 1}|y) = \frac{f(y|\text{Type 1})\Pr(\text{Type 1})}{f(y|\text{Type 1})\Pr(\text{Type 1}) + f(y|\text{Type 2})\Pr(\text{Type 2})}$$
$$= \frac{p \cdot \frac{1}{\sigma_1}\phi\left(\frac{y-\mu_1}{\sigma_1}\right)}{p \cdot \frac{1}{\sigma_1}\phi\left(\frac{y-\mu_1}{\sigma_1}\right) + (1-p) \cdot \frac{1}{\sigma_2}\phi\left(\frac{y-\mu_2}{\sigma_2}\right)}$$

# STATA Code: Components of Log-Likelihood

> mu1,mu2,sig1,sig2,p: 
$$\hat{\mu_1}, \hat{\sigma_1}, \hat{\mu_2}, \hat{\sigma_2}, \hat{p}$$
  
> f1:  $f(y|\text{Type 1}) = \frac{1}{\sigma_1} \phi\left(\frac{y-\mu_1}{\sigma_1}\right)$   
> f2:  $f(y|\text{Type 2}) = \frac{1}{\sigma_2} \phi\left(\frac{y-\mu_2}{\sigma_2}\right)$   
> log1:  
 $\ln[f(y)] = \ln\left[p \cdot \frac{1}{\sigma_1} \phi\left(\frac{y-\mu_1}{\sigma_1}\right) + (1-p) \cdot \frac{1}{\sigma_2} \phi\left(\frac{y-\mu_2}{\sigma_2}\right)\right]$   
> postp1:  $\Pr(\text{Type 1})$   
> postp2:  $\Pr(\text{Type 2})$ 

### STATA Code: Components of Log-Likelihood

program drop \_all

\* LIKELIHOOD EVALUATION PROGRAM STARTS HERE:

program define mixture

args logl mu1 sig1 mu2 sig2 p tempvar f1 f2

Global Variable: y Local Variable: 'mu1', 'sig1',...

\* GENERATE TYPE-CONDITIONAL DENSITIES: quietly gen double 'f1'=(1/'sig1')\*normalden((y-'mu1')/'sig1') quietly gen double 'f2'=(1/'sig2')\*normalden((y-'mu2')/'sig2')

\* COMBINE TYPE-CONDITIONAL DENSITIES WITH MIXING PROPORTIONS TO GENERATE MARGINAL DENSITY \* THIS IS THE FUNCTION THAT NEEDS TO BE MAXIMISED WHEN SUMMED OVER THE SAMPLE: quietly replace 'logl'=ln('p'\*'f1'+(1-'p')\*'f2')

```
* GENERATE THE POSTERIOR TYPE PROBABILITIES, AND MAKE THEM AVAILABLE OUTSIDE THE PROGRAM:
quietly replace postp1='p'*'f1'/('p'*'f1'+(1-'p')*'f2')
quietly replace postp2=(1-'p')*'f2'/('p'*'f1'+(1-'p')*'f2')
quietly putmata postp1, replace
```

```
program drop _all
* LIKELIHOOD EVALUATION PROGRAM STARTS HERE:
program define mixture
args logl mu1 sig1 mu2 sig2 p
tempvar f1 f2
```

```
* GENERATE TYPE-CONDITIONAL DENSITIES:
quietly gen double 'f1'=(1/'sig1')*normalden((y-'mu1')/'sig1')
quietly gen double 'f2'=(1/'sig2')*normalden((y-'mu2')/'sig2')
```

\* COMBINE TYPE-CONDITIONAL DENSITIES WITH MIXING PROPORTIONS TO GENERATE MARGINAL DENSITY \* THIS IS THE FUNCTION THAT NEEDS TO BE MAXIMISED WHEN SUMMED OVER THE SAMPLE: quietly replace 'logl'=ln('p'\*'f1'+(1-'p')\*'f2')

\* GENERATE THE POSTERIOR TYPE PROBABILITIES, AND MAKE THEM AVAILABLE OUTSIDE THE PROGRAM: quietly replace postp1='p'\*'f1'/('p'\*'f1'+(1-'p')\*'f2') quietly replace postp2=(1-'p')\*'f2'/('p'\*'f1'+(1-'p')\*'f2')

quietly putmata postp1, replace quietly putmata postp2, replace

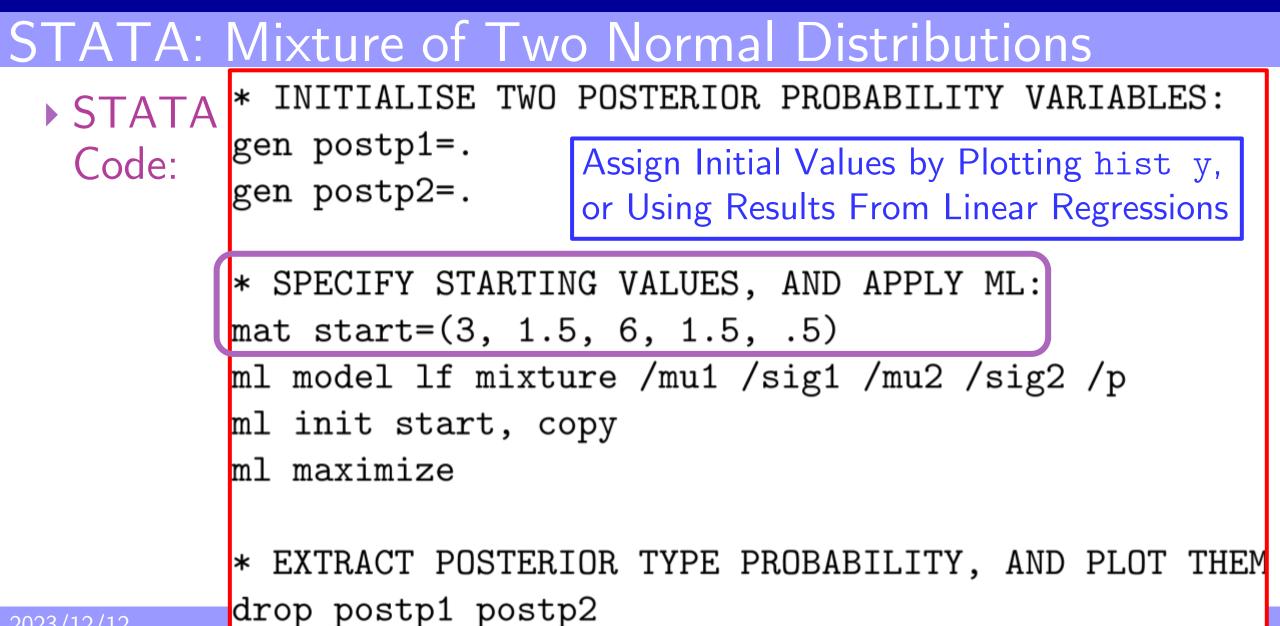
Save postp1, postp2 with STATA mata command putmata for later use

\* END OF LIKELIHOOD EVALUATION PROGRAM

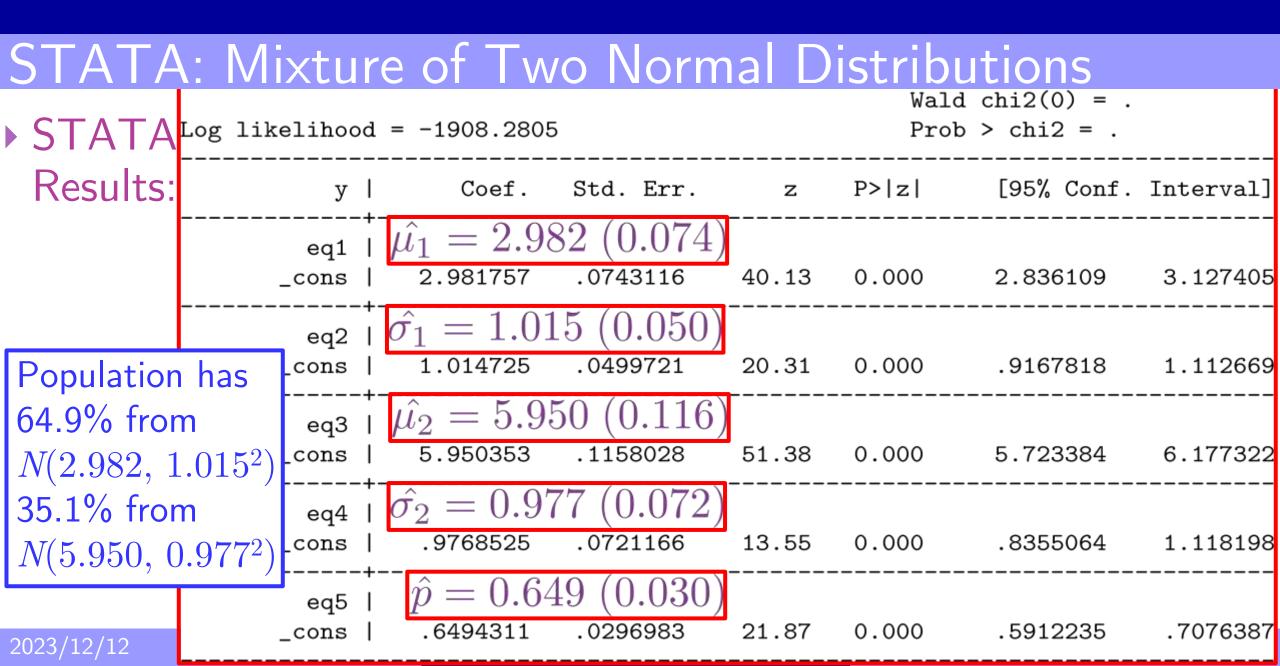
\* READ DATA:

end

use mixture\_sim, clear

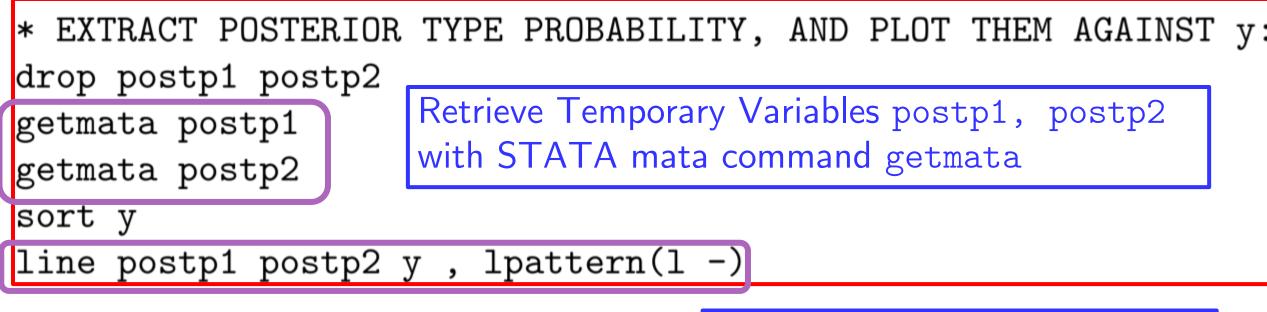


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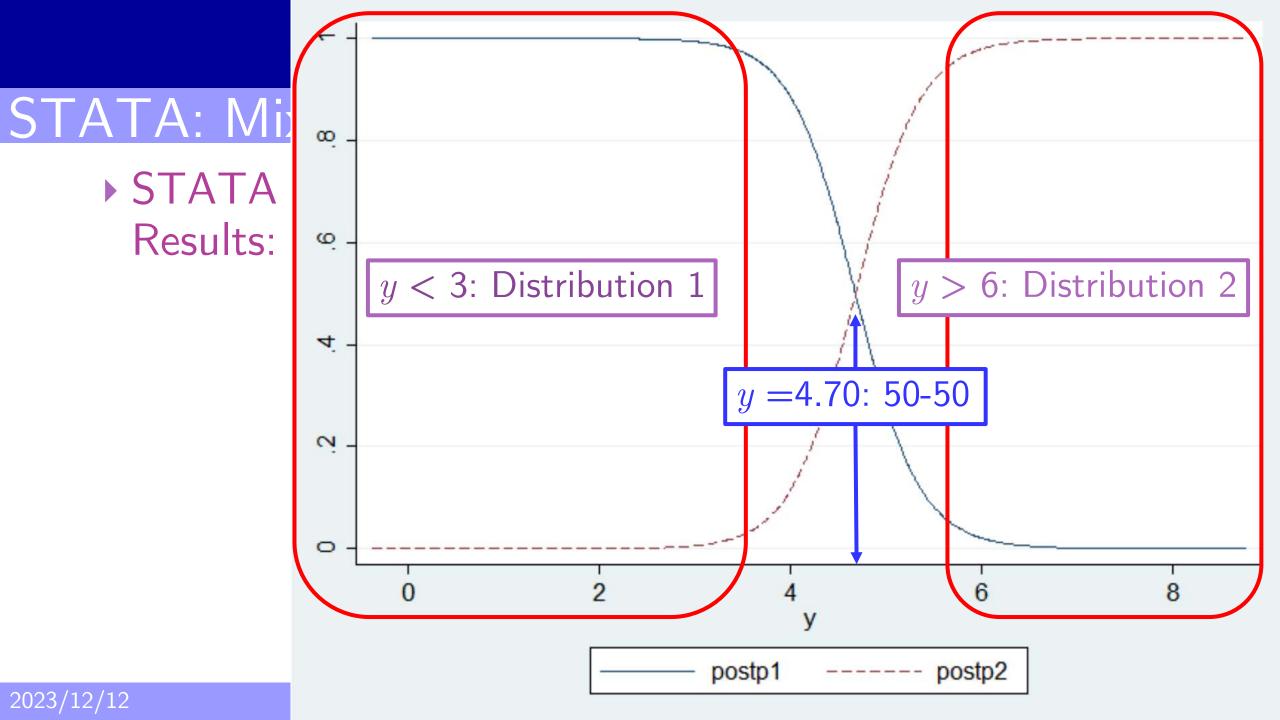
## STATA: Mixture of Two Normal Distributions

### STATA Code:

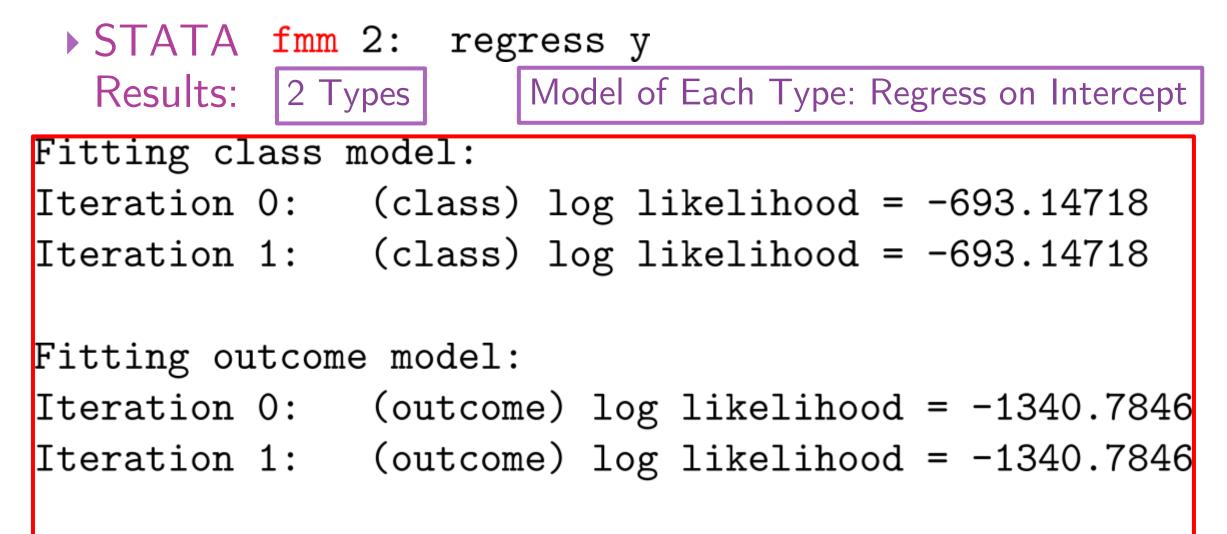


Plot Posterior Probability vs. y





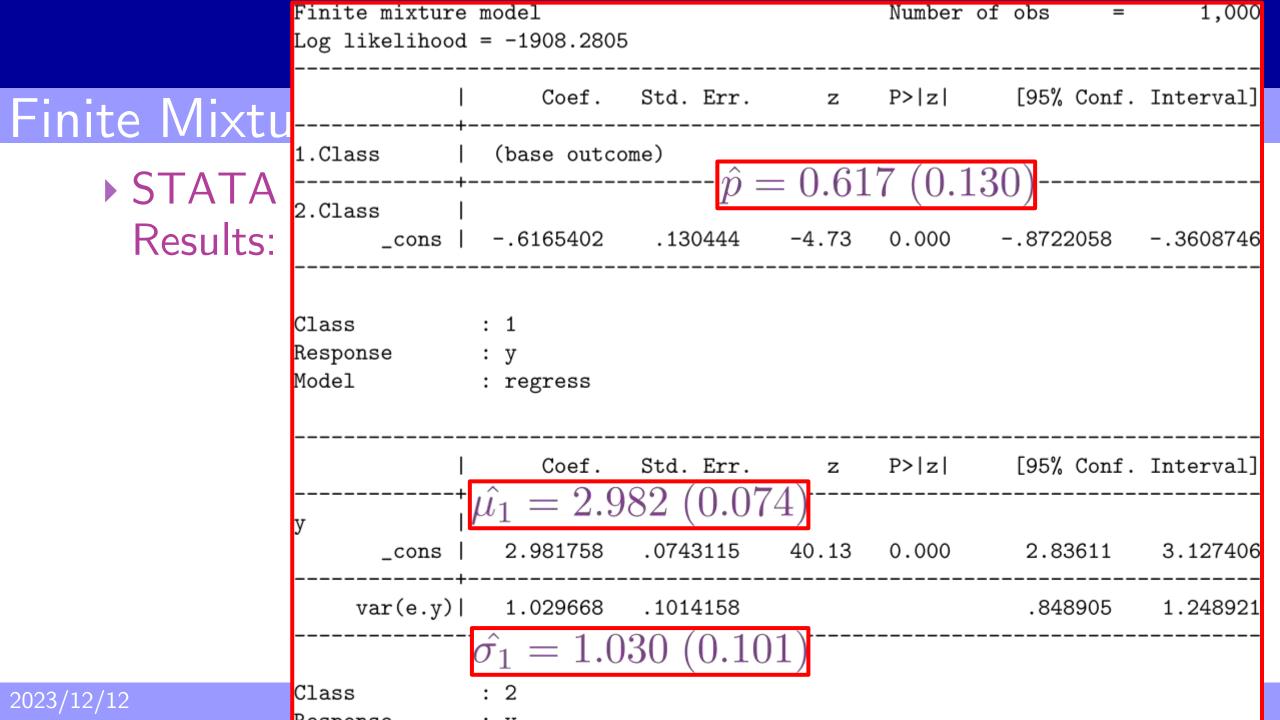
### Finite Mixture Model STATA Command: fmm

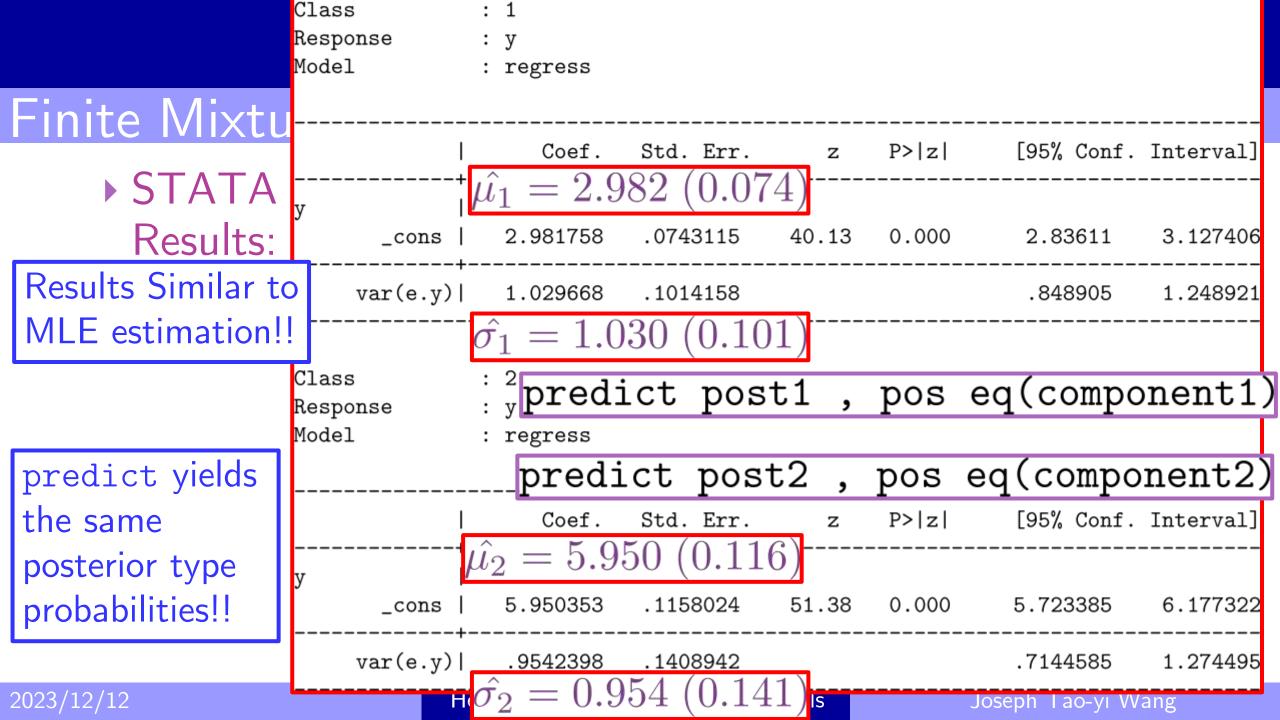


<sup>2023</sup>/Refining starting values:

	Refining s	starti	ng va	alues	3:			
	Iteration	0:	(EM)	log	likelihood	=	-2114.	989
Finite Mixtı	Iteration	1:	(EM)	log	likelihood	=	-2144.	1684
► STATA	Iteration	2:	(EM)	log	likelihood	=	-2155.	951
D L	Iteration	3:	(EM)	log	likelihood	=	-2159.	9264
Results:	Iteration Iteration	4:	(EM)	log	likelihood	=	-2159.	9464
	Iteration	5:	(EM)	log	likelihood	=	-2157.	8613
	Iteration	6:	(EM)	log	likelihood	=	-2154.	6472
	Iteration	7:	(EM)	log	likelihood	=	-2150.	8481
	Iteration	8:	(EM)	log	likelihood	=	-2146.	7758
	Iteration	9:	(EM)	log	likelihood	=	-2142.	6116
	Iteration	10:	(EM)	log	likelihood	=	-2138.	4622
	Iteration	11:	(EM)	log	likelihood	=	-2134.	3904
	Iteration	12:	(EM)	log	likelihood	=	-2130.	4335
2023/12/12	Iteration	13:	(EM)	log	likelihood	=	-2126.	6137
	<del>.</del>			-				~ ^ ^ ^

	Iteration	14:	(EM)	log	likelihood	= -212	2.9441
	Iteration		(EM)	log	likelihood	= -211	9.432
Finite Mixtu	Iteration	16:	(EM)	log	likelihood	= -211	6.0816
STATA Results:	Iteration	17:	(EM)	log	likelihood	= -211	2.8942
	Iteration	18:	(EM)	log	likelihood	= -210	9.8699
	Iteration	19:	(EM)	log	likelihood	= -210	7.0071
	Iteration	20:	(EM)	log	likelihood	= -210	4.3034
	Note: EM a	lgori	thm :	reach	ned maximum	iterat	ions.
	Fitting fu	ll mc	del:				
	Iteration	0:	log 1	likel	lihood = −1	909.813	7
	Iteration	1:	log :	likel	lihood = −1	908.403	1
	Iteration	2:	log 1	likel	lihood = −1	908.281	1
	Iteration	3:	log 1	likel	lihood = −1	908.280	5
2023/12/12	Iteration	4:	log 1	likel	lihood = −1	908.280	5





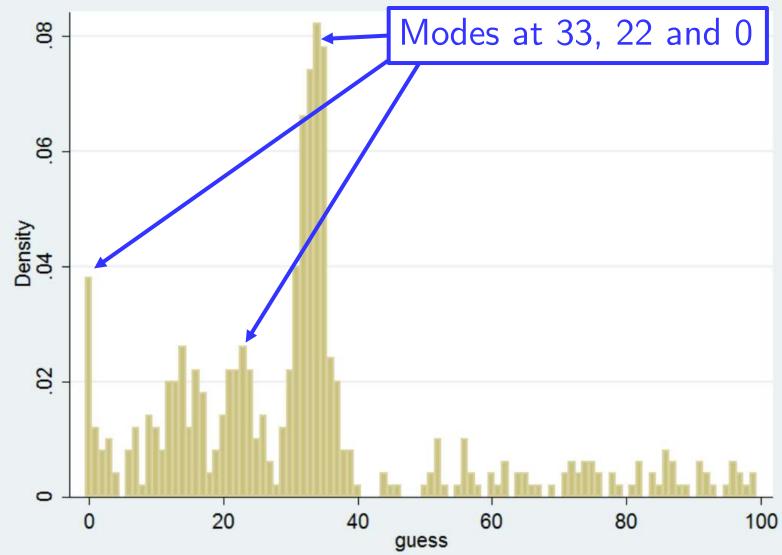
# Part II: A Level-*k* Model For The Beauty Contest Game 第二部分: 選美預測賽局的多層次認知模型

#### Joseph Tao-yi Wang (王道一) Experimetrics Lecture 6 (實驗計量第六講)

#### The *p*-Beauty Contest Game: Nagel (1995)

Het

- Choose a whole number in 0-100
- Number Closest to "p=2/3 of the Average" wins
  - Simulated Data of N=500 Players: beauty\_sim.dta



### <u>A Level-k Model For the Beauty Contest Game</u>

- Level-0 Reasoners Choose Randomly from Unif[0,100]
- Level-1 Believe Others are Level-0 and Choose 33
  Mean Guess = 50 and 50 x (2/3) = 33.333
- Level-2 Believe Others are Level-1 and Choose 22
  Mean Guess = 33 and 33 x (2/3) = 22
- Level-3 Believe Others are Level-2 and Choose 15
  - Mean Guess = 22 and  $22 \times (2/3) = 14.667$
- ▶ Level-4 Believe Others are Level-3 and Choose 10, etc.

## <u>A Level-k Model For the Beauty Contest Game</u>

- - ▶ All Guess 0 and Have Equal Chance to Win
- Same as Nash Equilibrium!
  - But real subjects do NOT play Nash (at least initially)
- ► To Estimate the Level-k Model:
  - Assume the Maximum Level = J
  - Let Level-J = naive-Nash (Choose Nash)
  - Let Level-0 choose randomly from uniform distribution

### Estimating the Level-k Model

- Level-j Chooses:  $y|_{\text{Type } j} = y_j^* + \epsilon, \ \epsilon \sim N(0, \sigma^2)$ 
  - Where  $y_j^* = \text{best guess of Type } j (j = 1, ..., J)$
- Conditional Density Functions:
  - Level-0:  $f(y|L_0) = 1/100, \ 0 \le y \le 100$

Level-j: 
$$f(y|L_j) = \frac{1}{\sigma} \phi\left(\frac{y - y_j^*}{\sigma}\right), \ 0 \le y \le 100 \ (j = 1, ..., J)$$

program define beauty\_mixture args lnf p1 p2 p3 p4 p5 sig tempvar f0 f1 f2 f3 f4 f5 l

quietly{

- J = 5
  - STATA: Maximized Log-Likelihood
- Best Guesses:
  - $y_1^* = 33.5$
  - $y_2^* = 22.4$
  - $y_3^* = 15.0$
  - $y_{4}^{*} = 10.1$

gen double 'f0'=0.01 gen double 'f1'=(1/'sig')\*normalden((y-33.5)/'sig') gen double 'f2'=(1/'sig')\*normalden((y-22.4)/'sig') gen double 'f3'=(1/'sig')\*normalden((y-15.0)/'sig') gen double 'f4'=(1/'sig')\*normalden((y-10.1)/'sig') gen double 'f5'=(1/'sig')\*normalden((y-0)/'sig')

gen double 'l'=(1-'p1'-'p2'-'p3'-'p4'-'p5')\*'f0' /// +'p1'\*'f1'+'p2'\*'f2'+'p3'\*'f3'+'p4'\*'f4'+'p5'\*'f5'

replace postp1='p1'\*'f1'/'l' replace postp2='p2'\*'f2'/'l' •  $y_5^* = 0$  (Naïve Nash) replace postp3='p3'\*'f3'/'1' replace postp4='p4'\*'f4'/'1' Heterogreplace postp5='p5'\*'f5'/'l'

replace 'lnf'=ln((1-'p1'-'p2'-'p3'-'p4'-'p5')\*'f0' /// +'p1'\*'f1'+'p2'\*'f2'+'p3'\*'f3'+'p4'\*'f4'+'p5'\*'f5')

putmata postp0, replace putmata postp1, replace putmata postp2, replace putmata postp3, replace putmata postp4, replace putmata postp5, replace

end

gen postp0=.

gen postp1=.

Heteroggen postp5=.

Best Guesses:

Estimating the level

STATA: Maximized

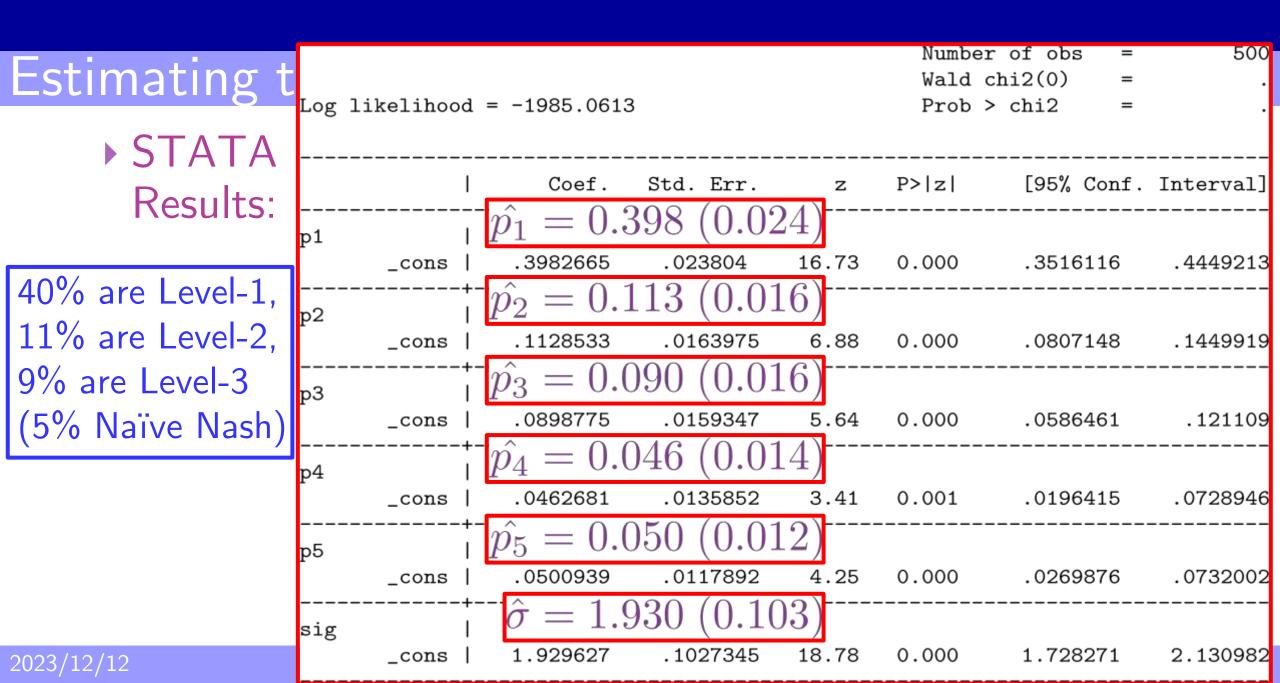
Log-Likelihood

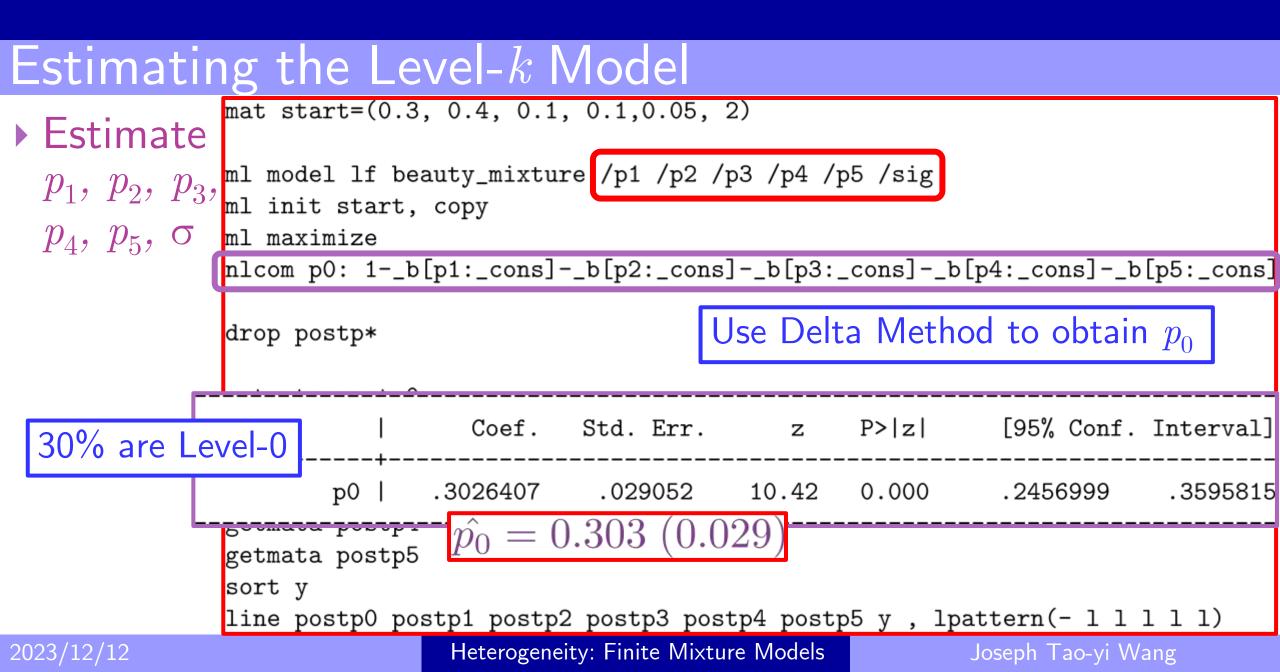
J = 5

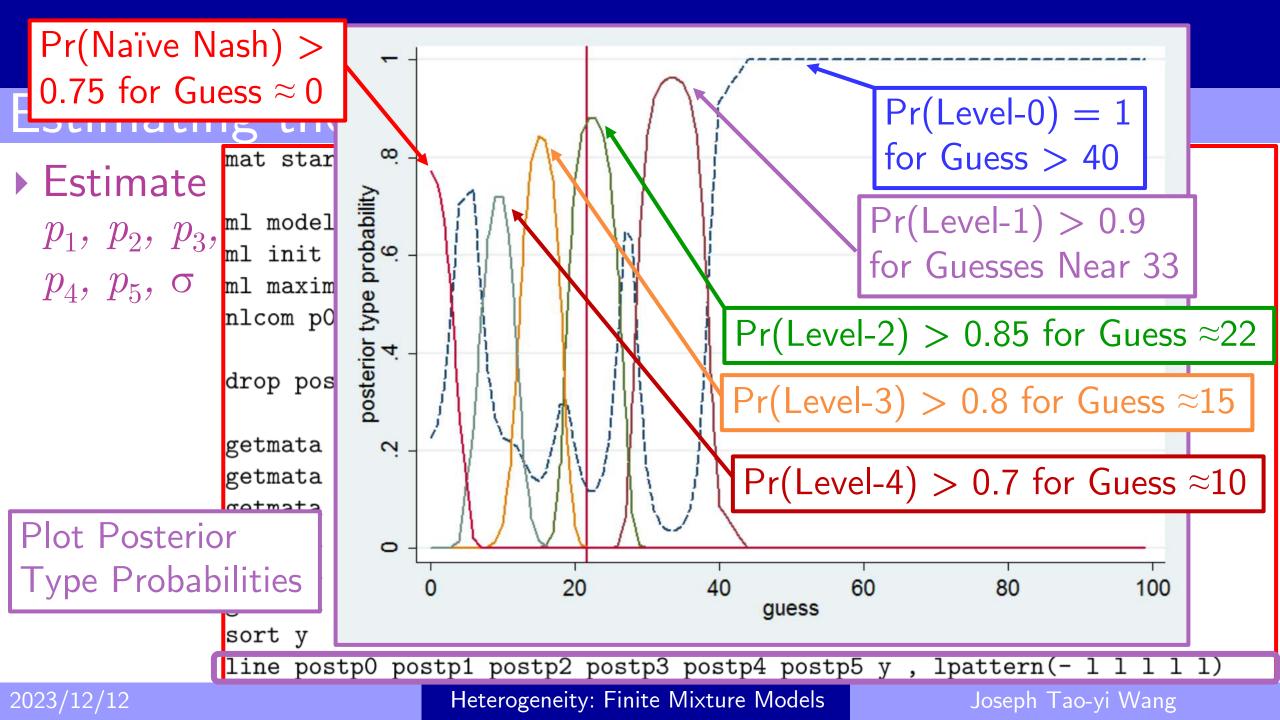
- $y_1^* = 33.5$
- $y_2^* = 22.4$
- $y_3^* = 15.0$
- $y_{4}^{*} = 10.1$
- gen postp2=. •  $y_5^* = 0$  (Naïve Nash)<sup>gen postp3=.</sup> gen postp4=.

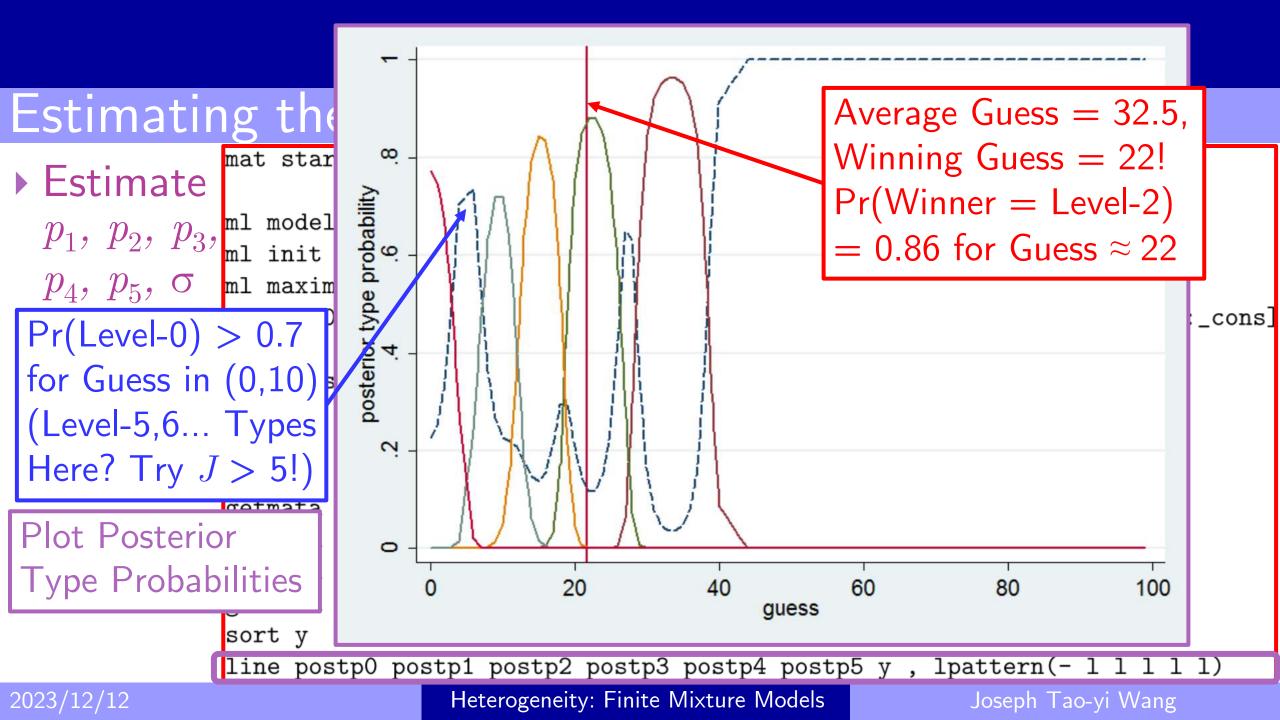
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Estimatir	ng the Level- $k$ Model
Estimate	mat start=(0.3, 0.4, 0.1, 0.1,0.05, 2)
	ml model lf beauty_mixture /p1 /p2 /p3 /p4 /p5 /sig ml init start, copy ml maximize
	nlcom p0: 1b[p1:_cons]b[p2:_cons]b[p3:_cons]b[p4:_cons]b[p5:_cons]
	drop postp*
	getmata postp0
	getmata postp1 getmata postp2
	getmata postp3 getmata postp4
	getmata postp5
	sort y line postp0 postp1 postp2 postp3 postp4 postp5 y , lpattern(- l l l l l)
0000/10/10	









# Part III: A Public Goods Game Experiment 第三部分: 公共財自願捐獻賽局實驗

#### Joseph Tao-yi Wang (王道一) Experimetrics Lecture 6 (實驗計量第六講)



# Public Goods Game Experiment

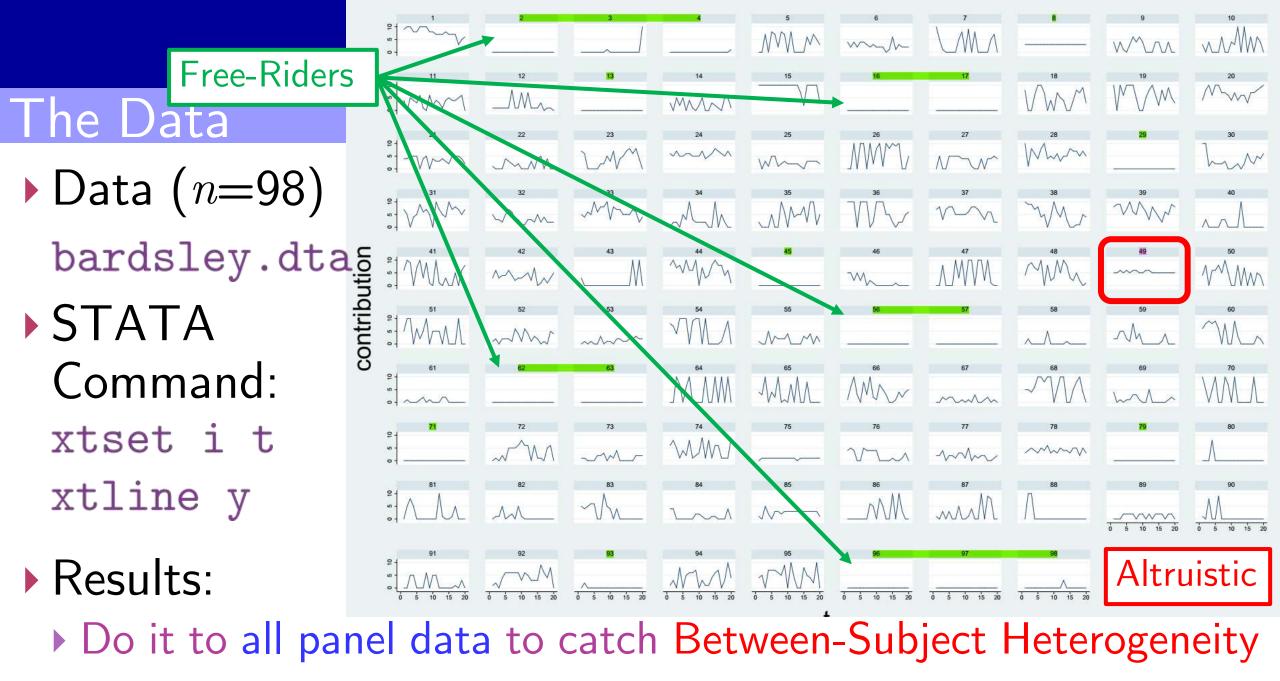
- ▶ n (= 7) Subjects per group with endowment  $e_i (= 10)$ 
  - Contribute to Public Account (or keep in Private Account)
  - Public account multiplied by k, but divided equally between all n members (MPCR = k/n)
- $\blacktriangleright$  Doubly Censored Data: Contribute between 0 and  $e_i$ 
  - ► Use Two-Limit Tobit Model (Nelson, 1976)
- Unique Nash: Zero Contribution
  - Experimental Data: Some positive contributions
  - Bardsley (2000): Uncover Motivations Behind Them

# Bardsley (2000): Why Contribution Decreases?

- 1. Learning to be Rational (learn incentive structure)
- 2. Social Learning (learn about others' behavior)
- Bardsley (2000): Conditional Information Lottery (CIL)
  - Play 1 Real Round mixed with 19 Fake Rounds against Computer, but only pay the real round
    - Subjects treat each round as real, but past rounds are not informative: They are fake if this round is real!
- Bardsley (2000): Take Turns to Contribute
  - See Previous Contributions Before Contributing

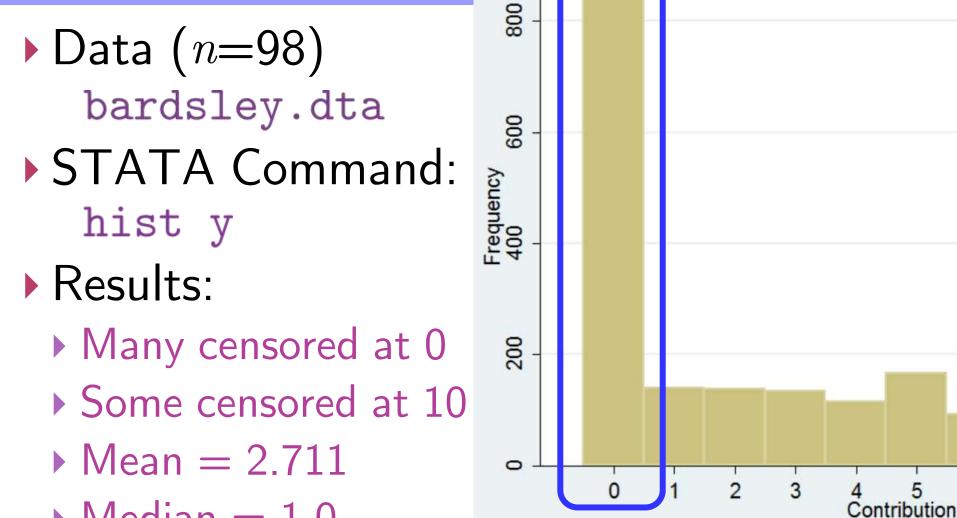
# Bardsley (2000): Take Turns to Contribute

- See Previous Contributions Before Contributing
- Use Mixture Model to Address Different Motivations:
- 1. Reciprocator (Depends on Previous Contributions)
  - Contributes if Median of Previous Contribution is High
- 2. Strategist (Depends on Position in Sequence)
  - Contributes to Induce Later Contributions
- 3. Free-Rider
  - Contributes 0 Regardless



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# The Data



► Median = 1.0

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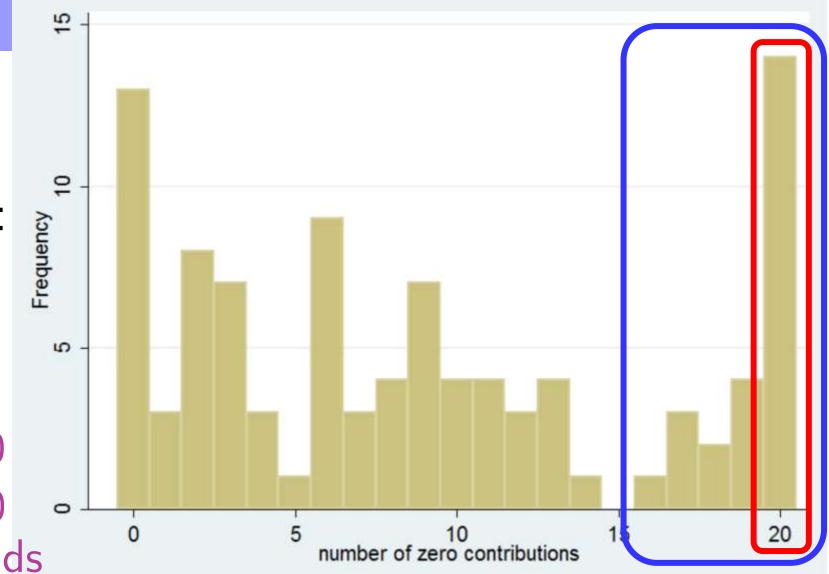
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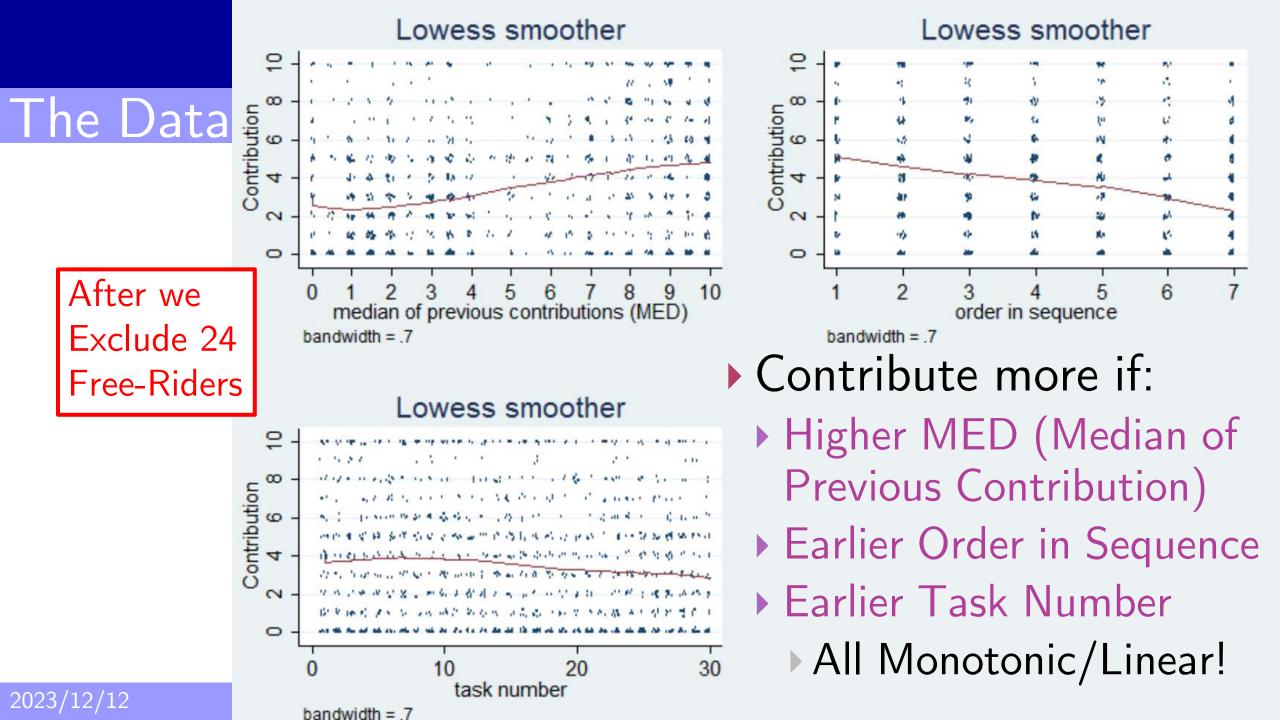
6

# <u>The Data</u>

- Data (n=98) bardsley.dta
  STATA Command: hist y=0 ?
- Results:
  - Identify Free Riders
  - ► 14.3% always give 0
  - 24.5% mostly give 0 in 16 out of 20 rounds



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- Bardsley and Moffatt (2007)
- $\blacktriangleright$  Observe n Subjects for T tasks
  - $\blacktriangleright$  Either Reciprocator, Strategist and Free-Rider for all T tasks
- Subject i contributes  $y_{it}$  in task t between 0 and 10
- 2-Limit Tobit Model for Reciprocator and Strategist:

 $y_{it} = \begin{cases} 0 & \text{if } y_{it}^* \leq 0 \quad (\text{Regime 1: No Contribution At All}) \\ y_{it}^* & \text{if } 0 < y_{it}^* < 10 (\text{Regime 2: Contribute b/w 0-10}) \\ 10 & \text{if } y_{it}^* \geq 10 \quad (\text{Regime 3: Full Contribution of 10}) \\ \hline \text{Desired} \end{cases}$ 

#### Finite Mixture 2-Limit Tobit Model with Tremble • Desired Contribution of Subjects i = 1 - n in tasks t = 1 - T are Reciprocator (rec) Median of Previous Contributions $y_{it}^* = \beta_{10} + \beta_{11}MED_{it} + \beta_{13}(TSK_{it} - 1) + \epsilon_{it,rec}$ >0 for Reciprocity <0: Learning $\epsilon_{it,rec} \sim N(0, \sigma_1^2)$ Desired Strategist (*str*) Decision Order Minus 1 Task Number (1-30) $y_{it}^* = \beta_{20} + \beta_{22}(ORD_{it} - 1) + \beta_{23}(TSK_{it} - 1) + \epsilon_{it,str}$ E(Contribution | <0 for Strategic Behavior $\epsilon_{it,str} \sim N(0,\sigma_2^2)$ Task 1, Order 1) Free-Rider (fr): None $y_{it} = 0$

Heterogeneity: Finite Mixture Models

- Prior Expectation of Others' Contribution
  - Set MED = 8.00 if ORD = 1 (trial-and-error to max. log-L)
- Mistakes (Moffatt and Peters, 2001): Tremble  $\omega$ 
  - Decreasing magnitude over time  $\omega_{it} = \omega_0 \exp \left[\omega_1 (TSK_{it} 1)\right]$
  - $\blacktriangleright$  Initial tremble probability  $\omega_0~$  vs. rate of decay  $\omega_1 < 0$
- Regime 1 (y = 0)
- Regime 2 (0 < y < 10)
- Regime 3 (y = 10)

# Regime 1 (y = 0): Tremble: 0-10 with Equal Chance

 $\Pr(y_{it} = 0 | i = \operatorname{rec}) = \frac{1}{(1 - \omega_{it})} \Phi\left(\frac{-\beta_{10} - \beta_{11}MED_{it} - \beta_{13}(TSK_{it} - 1)}{\sigma_1}\right) + \frac{\omega_{it}}{11}$   $\Pr(y_{it} = 0 | i = \operatorname{str}) = \frac{(-\beta_{20} - \beta_{22}(ORD_{it} - 1) - \beta_{23}(TSK_{it} - 1))}{\omega_{it}}$ 

$$(1 - \omega_{it})\Phi\left(\frac{-\beta_{20} - \beta_{22}(ORD_{it} - 1) - \beta_{23}(ISR_{it} - 1)}{\sigma_2}\right) + \frac{\omega_{it}}{11}$$
  

$$\Pr(y_{it} = 0|i = \text{fr}) = 1 - \frac{10\omega_{it}}{11}$$

# Finite Mixture 2-Limit Tobit Model with Tremble • Regime 2 (0 < y < 10): **Tremble**: Uniform[-0.5, 10.5] $f(y_{it}|i = \text{rec}) =$ $(1-\omega_{it})\frac{1}{\sigma_1}\Phi\left(\frac{y_{it}-\beta_{10}-\beta_{11}MED_{it}-\beta_{13}(TSK_{it}-1)}{\sigma_1}\right)+\frac{\omega_{it}}{11}$ • $f(y_{it}|i = \operatorname{str}) =$ $(1 - \omega_{it})\frac{1}{\sigma_2}\Phi\left(\frac{y_{it} - \beta_{20} - \beta_{22}(ORD_{it} - 1) - \beta_{23}(TSK_{it} - 1)}{\sigma_2}\right) + \frac{\omega_{it}}{11}$

• 
$$f(y_{it}|i=\mathrm{fr}) = \frac{\omega_{it}}{11}$$

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► Regime 3 
$$(y = 10)$$
:  
►  $Pr(y_{it} = 10|i = rec) =$   
 $(1 - \omega_{it}) \left[ 1 - \Phi \left( \frac{10 - \beta_{10} - \beta_{11}MED_{it} - \beta_{13}(TSK_{it} - 1)}{\sigma_1} \right) \right] + \frac{\omega_{it}}{11}$   
►  $Pr(y_{it} = 10|i = str) =$ 

$$(1 - \omega_{it}) \left[ 1 - \Phi \left( \frac{10 - \beta_{20} - \beta_{22}(ORD_{it} - 1) - \beta_{23}(TSK_{it} - 1)}{\sigma_2} \right) \right] + \frac{\omega_{it}}{11}$$
  
 
$$\Pr(y_{it} = 10 | i = \text{fr}) = \frac{\omega_{it}}{11}$$

# Likelihood Function is L<sub>i</sub>

$$= p_{\rm rec} \prod_{t=1}^{T} \Pr(y_{it} = 0|\text{rec})^{I_{y_{it}=0}} f(y_{it}|\text{rec})^{I_0 < y_{it} < 10} \Pr(y_{it} = 10|\text{rec})^{I_{y_{it}=10}} + p_{\rm str} \prod_{t=1}^{T} \Pr(y_{it} = 0|\text{str})^{I_{y_{it}=0}} f(y_{it}|\text{str})^{I_0 < y_{it} < 10} \Pr(y_{it} = 10|\text{str})^{I_{y_{it}=10}} + p_{\rm fr} \prod_{t=1}^{T} \Pr(y_{it} = 0|\text{fr})^{I_{y_{it}=0}} f(y_{it}|\text{fr})^{I_0 < y_{it} < 10} \Pr(y_{it} = 10|\text{fr})^{I_{y_{it}=10}}$$

 $\hat{\beta}_{10}, \dots, \hat{\beta}_{23}, \hat{\sigma}_1, \hat{\sigma}_2; \hat{\omega}_0, \hat{\omega}_1; \hat{p}_{\text{rect}}, \hat{p}_{\text{str}}, \hat{p}_{\text{fr}} \text{ maximize} \log L = \sum \log(L_i)$ □(Sample Log-Likelihood) i=12023/12/12

Heterogeneity: Finite Mixture Models

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# STATA Code: Components of Log-Likelihood

- p1\_1,p2\_1,p3\_1:Pr(y = 0|rec), Pr(y = 0|str), Pr(y = 0|fr)
  p1\_2,p2\_2,p3\_2:f(y|rec), f(y|str), f(y|fr), 0 < y < 10</li>
- ▶ p1\_3,p2\_3,p3\_3:Pr(y = 10|rec), Pr(y = 10|str), Pr(y = 10|fr)
- ▶ p1:

$$\Pr(y_{it} = 0 | \operatorname{rec})^{I_{y_{it}}=0} f(y_{it} | \operatorname{rec})^{I_0 < y_{it} < 10} \Pr(y_{it} = 10 | \operatorname{rec})^{I_{y_{it}}=10}$$
p2:

$$\Pr(y_{it} = 0 | \text{str})^{I_{y_{it}}=0} f(y_{it} | \text{str})^{I_0 < y_{it} < 10} \Pr(y_{it} = 10 | \text{str})^{I_{y_{it}}=10}$$
  
• p3:

$$\Pr(y_{it} = 0|\mathrm{fr})^{I_{y_{it}}=0} f(y_{it}|\mathrm{fr})^{I_{0 < y_{it} < 10}} \Pr(y_{it} = 10|\mathrm{fr})^{I_{y_{it}}=10}$$

# STATA Code: Components of Log-Likelihood

# $\prod_{t=1}^{t} \Pr(y_{it} = 0|\operatorname{rec})^{I_{y_{it}}=0} f(y_{it}|\operatorname{rec})^{I_{0 < y_{it} < 10}} \Pr(y_{it} = 10|\operatorname{rec})^{I_{y_{it}}=10}$

# ▶ pp2:<sub>T</sub> $\prod_{t=1}^{T} \Pr(y_{it} = 0 | \operatorname{str})^{I_{y_{it}=0}} f(y_{it} | \operatorname{str})^{I_{0 < y_{it} < 10}} \Pr(y_{it} = 10 | \operatorname{str})^{I_{y_{it}=10}}$

▶ pp3:<sub>T</sub>  $\prod_{t=1}^{T} \Pr(y_{it} = 0 | \text{fr})^{I_{y_{it}=0}} f(y_{it} | \text{fr})^{I_{0 < y_{it} < 10}} \Pr(y_{it} = 10 | \text{fr})^{I_{y_{it}=10}}$ 

# STATA Code: Components of Log-Likelihood

- theta1:  $\beta_{10}, \beta_{11}, \beta_{13}$
- ▶ theta2:  $\beta_{20}, \beta_{22}, \beta_{23}$
- imes sig1, sig2, w0, w1, w:  $\sigma_1, \sigma_2, \omega_0, \omega_1, \omega$
- $p_{rec}, p_{str}, p_{fr}: p_{rect}, p_{str}, p_{fr}$
- $pp, lnpp: L_i, LogL = \sum_{i=1}^n \log(L_i)$
- ▶ postp1:  $Pr(i = rec|y_{i1}, \ldots, y_{iT})$
- ▶ postp2:  $\Pr(i = \operatorname{str}|y_{i1}, \ldots, y_{iT})$
- ▶ postp3:  $Pr(i = fr|y_{i1}, \ldots, y_{iT})$

\* ESTIMATION OF MIXTURE MODEL FOR BARDSLEY DATA

prog drop \_all

\* LIKELIHOOD EVALUATION PROGRAM STARTS HERE:

program define pg\_mixture args todo b lnpp tempvar p1\_1 p2\_1 p3\_1 p1\_2 p2\_2 p3\_2 p1\_3 p2\_3 p3\_3 p1 p2 p3 pp1 pp2 pp3 pp w

tempname theta1 theta2 sig1 sig2 w0 w1 p\_rec p\_str

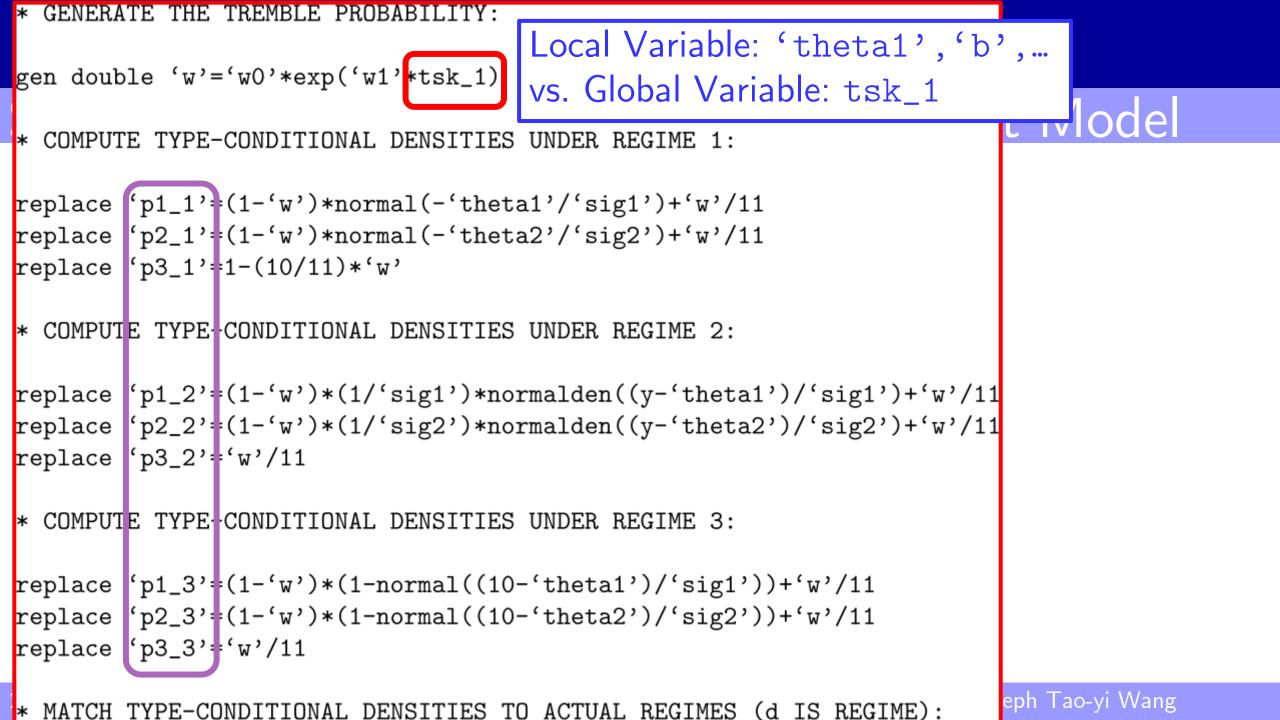
\* ASSIGN PARAMETER NAMES TO THE ELEMENTS OF THE PARAMETER VECTOR b:

mleval 'theta1' = 'b' eq(1)
mleval 'theta2' = 'b' eq(2)
mleval 'sig1' = 'b', eq(3) scalar
mleval 'sig2'='b', eq(4) scalar
mleval 'w0'='b', eq(5) scalar
mleval 'w1'='b', eq(6) scalar
mleval 'p\_rec'='b', eq(7) scalar
mleval 'p\_str'='b', eq(8) scalar

Local Variable: 'theta1', 'b',... vs. Global Variable: tsk\_1 (below)

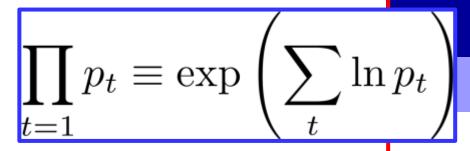


mieval 'p_rec'='b', $e_{I}(7)$ sca	alar		
<pre>mleval 'p_str'='b', eq(8) sca</pre>	alar		
quietly{		2-1 imi	t Tobit Model
* INITIALISE THE p* VARIABLES	C WITH MIGGING VALUES.		
$\star$ INTITALISE IIIE $P^{\star}$ VARIADEEC	S WIIII MISSING VALUES.		
gen double 'p1_1'=.			
gen double 'p2_1'=.			
gen double 'p3_1'=.			
gen double 'p1_2'=.			
gen double 'p2_2'=.			
gen double 'p3_2'=.			
gen double 'p1_3'=.			
gen double 'p2_3'=.	Local Variable: 't		
gen double 'p3_3'=.	vs. Global Variabl	e: tsk 1	(below)
gen double 'p1'=.			
gen double 'p2'=.			
gen double 'p3'=.			
gen double 'pp1'=.			
gen double 'pp2'=.			
gen double 'pp3'=.			
gen double 'pp'=.		ture Models	Joseph Tao-yi Wang



\* MATCH TYPE-CONDITIONAL DENSITIES TO ACTUAL REGIMES (d IS REGIME):

replace 'p1' = (d==1)\*'p1\_1'+(d==2)\*'p1\_2'+(d==3)\*'p1\_3'
replace 'p2' = (d==1)\*'p2\_1'+(d==2)\*'p2\_2'+(d==3)\*'p2\_3'
replace 'p3' = (d==1)\*'p3\_1'+(d==2)\*'p3\_2'+(d==3)\*'p3\_3'



\* FIND PRODUCT OF TYPE-CONDITIONAL DENSITIES FOR EACH SUBJECT:

by	i:	replace	<pre>'pp1'=exp(sum(ln(max('p1',</pre>	1e-12)	)))
by	i:	replace	<pre>'pp2'=exp(sum(ln(max('p2',</pre>	1e-12)	)))
by	i:	replace	<pre>'pp3'=exp(sum(ln(max('p3',</pre>	1e-12)	)))

Sum  $\ln(p_1)$  instead of product

Use "1e-12" if close to 0 to avoid negative infinity at ln(0)

\* COMBINE TYPE-CONDITIONAL DENSITIES TO OBTAIN MARGINAL DENSITY FOR EACH SUBJECT \* (ONLY REQUIRED IN FINAL ROW FOR EACH SUBJECT):

replace 'pp'= p\_rec'\*'pp1'+'p\_str'\*'pp2'+(1-'p\_rec'-'p\_str')\*'pp3'
replace 'pp'= if last~=1

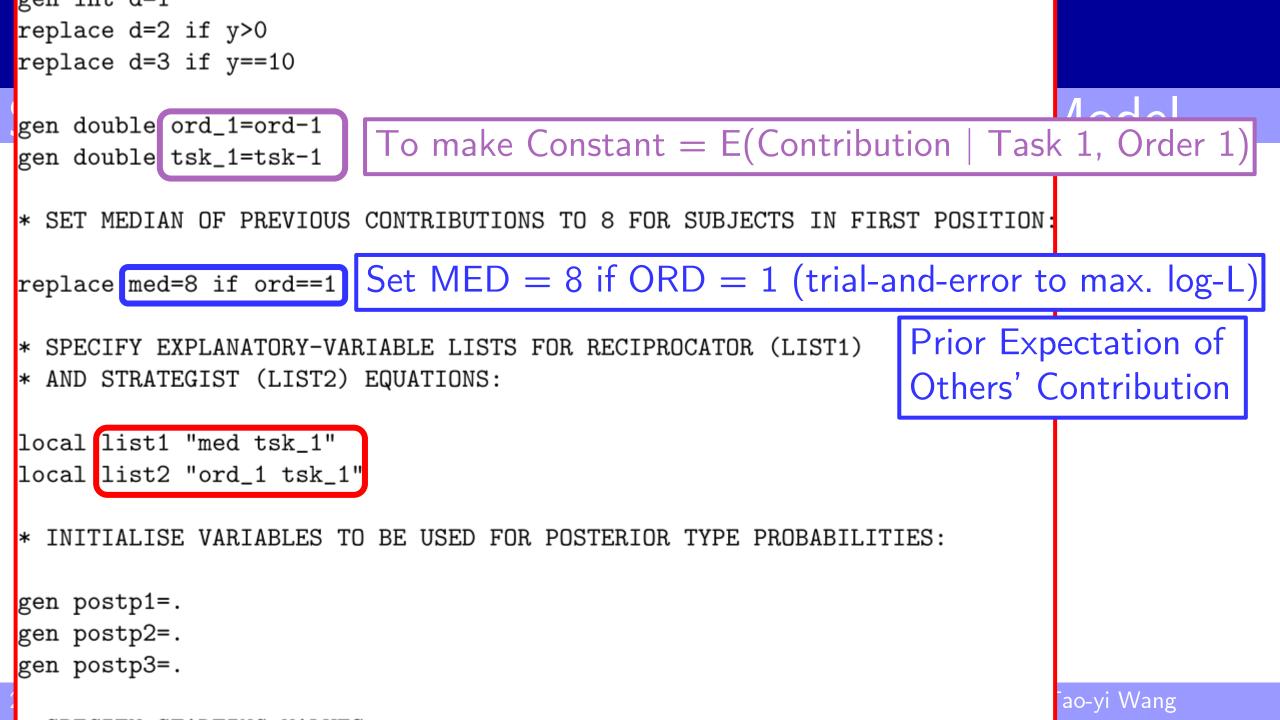
\* SPECIFY (LOG-LIKELIHOOD) FUNCTION WHOSE SUM OVER SUBJECTS IS TO BE MAXIMISED

```
mlsum 'lnpp'=_n('pp') if last==1
```

\* GENERATE POSTERIOR TYPE PROBABILITIES, AND MAKE THESE AVAILABLE OUTSIDE THE PROGRAM

\* GENERATE POSTERIOR TYPE PROBABILITIES, AND MAKE THESE AVAILABLE OUTSIDE THE PROGRAM

```
replace postp1='p_rec'*'pp1'/'pp'
replace postp2='p_str'*'pp2'/'pp'
replace postp3=(1-'p_rec'-'p_str')*'pp3'/'pp'
putmata postp1, replace
putmata postp2, replace
putmata postp3, replace
end
* END OF LOG-LIKELIHOOD EVALUATION PROGRAM
clear
set more off
* READ DATA
                       Data: bardsley.dta
use 'bardsley'
by i: gen last=_n==_N
gen int d=1
```



#### \* SPECIFY STARTING VALUES:

mat start=(0.57,-0.10,6.1,-0.93,-0.05,5.2,3.3,3.7,0.11,-0.05,0.26,0.49)

\* SPECIFY LIKELIHOOD EVALUATOR, PROGRAM, AND PARAMETER NAMES:

ml model d0 pg\_mixture (='list1') (='list2') /sig1 /sig2 /w0 /w1 /p1 /p2

<sup>ml init start, copy</sup> Cannot use lf since mixture model has non-linear log-L

Use D-Family: d0 requires only log-L

\* USE ML COMMAND TO MAXIMISE LOG-LIKELIHOOD, AND STORE RESULTS AS "WITH\_TREMBLE"

ml max, trace search(norescale)
est store with\_tremble

\* COMPUTE THIRD MIXING PROPORTION USING DELTA METHOD:

nlcom p3: 1-[p1]\_b[\_cons]-[p2]\_b[\_cons]

#### Derive p3 using the Delta Method!

(d1/d2 requires analytical derivatives of log-L)

\* EXTRACT POSTERIOR TYPE PROBABILITIES AND PLOT THEM AGAINST \* NUMBER OF ZERO CONTRIBUTIONS:

drop postp1 postp2 postp3

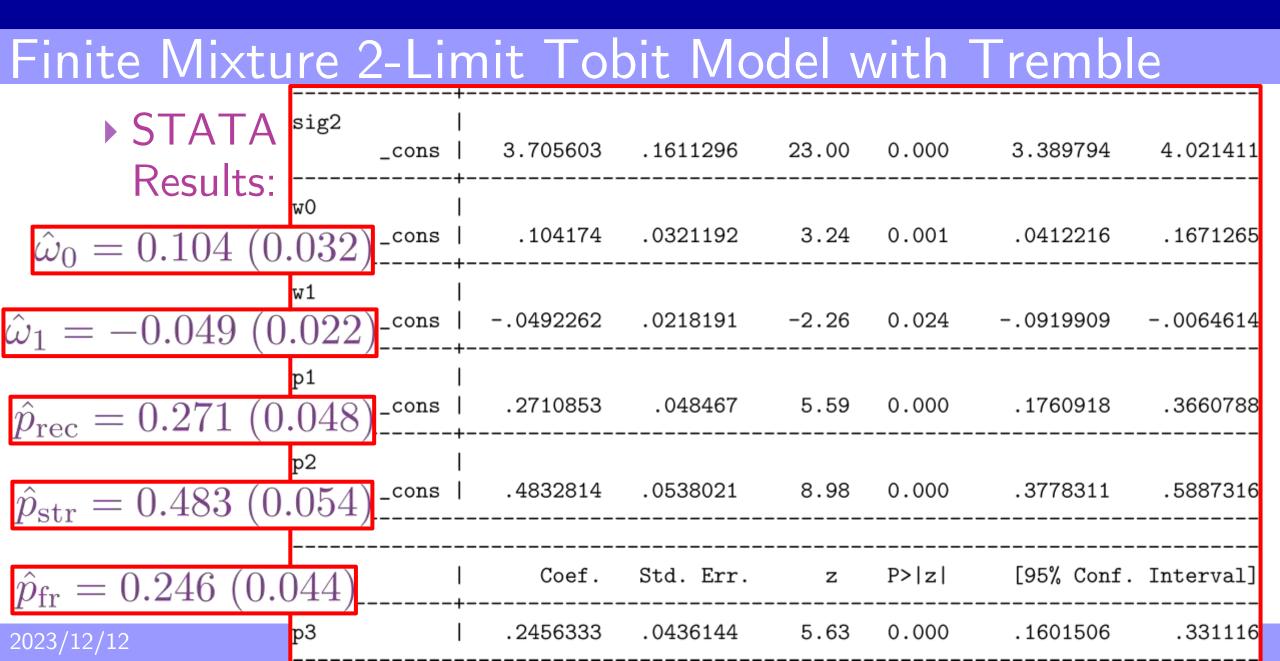
						Numbe	er of obs =	1960
Finite Mixtu							chi2(2) =	108.07
	.og lik	elihood	l = -3267.6884	1			> chi2 =	0.0000
		.0111000	0201.000	-		1100	01112	0.0000
► STATA _								
		1	Coef.	Std. Err.	z	P> z	[95% Conf.	Interval]
Results:		י +						
e	eq1							
		med	.598677	.0611812	9.79	0.000	.4787641	.7185899
$\beta_{11} = 0.599 \ (0.06)$	)1)	tsk_1	0961739	.0202229	-4.76	0.000	13581	0565379
$\hat{\beta}_{13} = -0.096 \ (0.0)$		_cons	4.004374	.4541832	8.82	0.000	3.114192	4.894557
$p_{13} = -0.050$ (0.0	020)	+	ô	1 00 1 (0	4 F 4)			
e	eq2	I	$\beta_{10} =$	4.004(0.	454)			
$\hat{B}_{1} = 0.064(0.0)$	000)	ord_1	9644643	.0823741	-11.71	0.000	-1.125915	803014
$\hat{\beta}_{22} = -0.964 \ (0.00)$	062)	tsk_1	0516766	.017189	-3.01	0.003	0853664	0179867
$\hat{\beta}_{23} = -0.052 \ (0.00)$	(017)	_cons	5.299353	.3828498	13.84	0.000	4.548981	6.049724
		+	· Â —	5.299(0.	383)			
s	sig1		$P_{20} =$	0.200 (0.	000)			
$\hat{\sigma}_1 = 3.442 \ (0.1)$	L67)	_cons	3.442241	.1674649	20.56	0.000	3.114016	3.770466
s	sig2							
$_{20}\hat{\sigma}_2 = 3.706~(0.1)$		_cons	3.705603	.1611296	23.00	0.000	3.389794	4.021411

Reciprocator (rec)
$$y_{it}^* = \beta_{10} + \beta_{11}MED_{it} + \beta_{13}(TSK_{it} - 1) + \epsilon_{it,rec}$$

$$E(y^*|MED, TSK) = 4.004 + 0.599MED - 0.096(TSK - 1)$$
Strategist (str)
$$>0 \& <1 \text{ for Biased Reciprocity} \quad <0: \text{ Learning}$$

$$y_{it}^* = \beta_{20} + \beta_{22}(ORD_{it} - 1) + \beta_{23}(TSK_{it} - 1) + \epsilon_{it,str}$$

$$E(y^*|ORD, TSK) = 5.299 - 0.964(ORD - 1) - 0.052(TSK - 1)$$
Slower than Reciprocators
$$Slower than Reciprocators$$

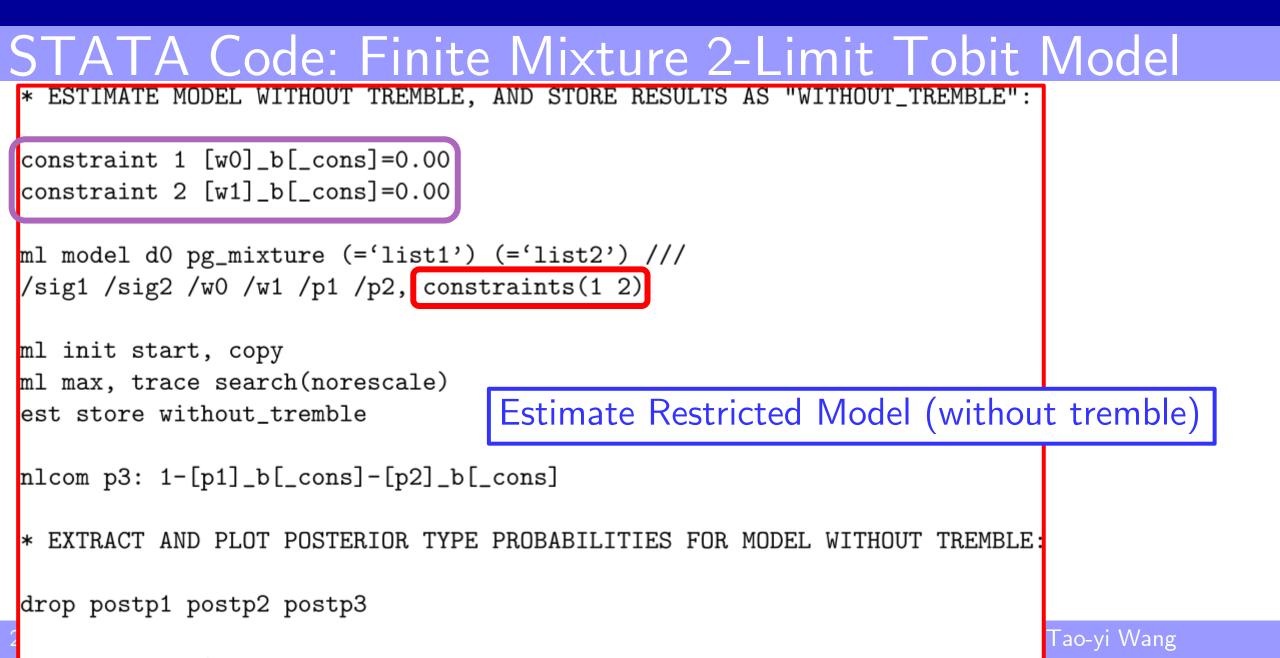


# <u>STATA Code: Finite Mixture 2-Limit Tobit Model</u>

```
* EXTRACT POSTERIOR TYPE PROBABILITIES AND PLOT THEM AGAINST
```

```
* NUMBER OF ZERO CONTRIBUTIONS:
```

```
drop postp1 postp2 postp3
getmata postp1
getmata postp2
getmata postp3
label variable postp1 "rec"
label variable postp2 "str"
label variable postp3 "fr"
                                  Plot posterior probabilities (with tremble)
by i: gen n_zero=sum(y==0)
scatter postp1 postp2 postp3 n_zero if last==1, title("with tremble") ///
ytitle("posterior probability") msymbol(x Dh Sh) jitter(3) saving(with, replace)
```



```
* EXTRACT AND PLOT POSTERIOR TYPE PROBABILITIES FOR MODEL WITHOUT TREMBLE:
```

```
drop postp1 postp2 postp3
```

getmata postp1 getmata postp2 getmata postp3

label variable postp1 "rec" label variable postp2 "str" label variable postp3 "fr"

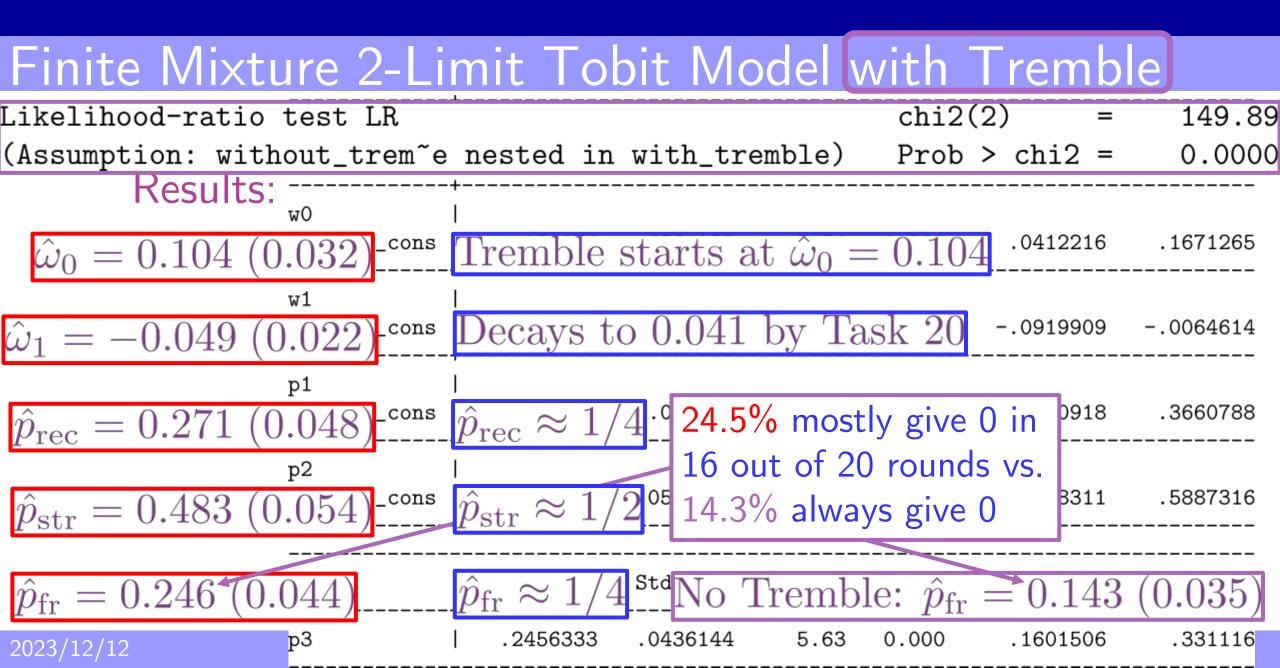
```
scatter postp1 postp2 postp3 n_zero if last==1, title("without tremble") ///
ytitle("posterior probability") msymbol(x Dh Sh) jitter(3) saving(without, replace)
```

\* CARRY OUT LIKELIHOOD RATIO TEST FOR PRESENCE OF TREMBLE:

Irtest with\_tremble without\_tremble Likelihood Ratio Test (with/without tremble)

\* COMBINE THE TWO POSTERIOR PROBABILITY PLOTS

gr combine with.gph without.gph



# Posterior Type Probabilities

$$\Pr(i = \operatorname{rec}|y_{i1}, \dots, y_{iT}) = \frac{p_{\operatorname{rec}}}{L_i} \prod_{t=1}^T \Pr(y_{it} = 0|\operatorname{rec})^{I_{y_{it}=0}} f(y_{it}|\operatorname{rec})^{I_0 < y_{it} < 10} \Pr(y_{it} = 10|\operatorname{rec})^{I_{y_{it}=10}}$$

$$\Pr(i = \operatorname{str}|y_{i1}, \dots, y_{iT}) = \frac{p_{\operatorname{str}}}{L_i} \prod_{t=1}^T \Pr(y_{it} = 0|\operatorname{str})^{I_{y_{it}}=0} f(y_{it}|\operatorname{str})^{I_0 < y_{it} < 10} \Pr(y_{it} = 10|\operatorname{str})^{I_{y_{it}}=10}$$

$$\Pr(i = \operatorname{fr}|y_{i1}, \dots, y_{iT}) = \frac{p_{\operatorname{fr}}}{L_i} \prod_{t=1}^T \Pr(y_{it} = 0 | \operatorname{fr})^{I_{y_{it}=0}} f(y_{it} | \operatorname{fr})^{I_0 < y_{it} < 10} \Pr(y_{it} = 10 | \operatorname{fr})^{I_{y_{it}=10}}$$

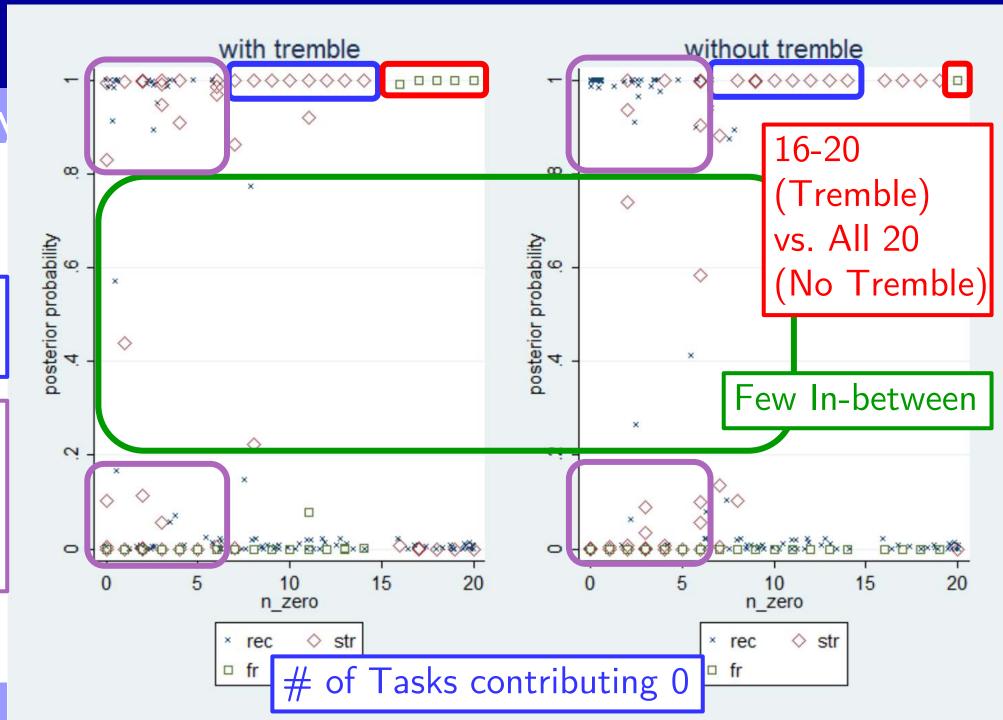
Posterior Ty

STATA Results:

6-14(or 6-19) are Strategists

0-5 are Mixture of Strategists and Reciprocators

2023/12/12



# Conclusion: Finite Mixture Model

- Mixture Model accounts for Types in the Population
  - Infinite Mixture Model = Random Coefficient Model

# How it Works?

- Economic Theory Predicts and Name Various Types
- Construct Parametric Model for Behavior of Each Type
- Estimated Using Population Data to Obtain:
  - Mixing Proportions and Parameters of Each Type
  - Individual Posterior Probability of being a Type

## Acknowledgment

## This presentation is based on

- ▶ Section 5.1-5 of the lecture notes of Experimetrics,
- Prepared for a mini-course taught by Peter G. Moffatt (UEA) at National Taiwan University in Spring 2019
  - ▶ We would like to thank 康柏賢 for his in-class presentations

