## Econometrics I

Midterm Exam. II
Dec. 13, 2002

1. (20\%) State with brief reasons whether the following statements are true, false, or uncertain.
(a) (5\%) With the fact that $\left(A^{-1}\right)^{\prime}=\left(A^{\prime}\right)^{-1}$, if $A$ is symmetric, then $A^{-1}$ is also symmetric.
(b) (5\%) If we drop an explanatory variable from a regression model, both $R^{2}$ and $\bar{R}^{2}$ (adjusted $R^{2}$ ) will decrease.
(c) (5\%) For a simple regression model without constant term, $Y_{i}=\beta X_{i}+u_{i}$, the OLS estimator of $\beta$ is $\frac{\bar{Y}}{\bar{X}}$, where $\bar{Y}$ and $\bar{X}$ are the average of $Y_{i}$ and $X_{i}$ respectively.
(d) (5\%) If $A$ and $B$ are two $m \times m$ matrices, therefore both $A B$ and $B A$ exist. Then $A B=B A$.
2. $(15 \%)$ The following regression is estimated as a production function:

$$
\ln Q=1.37+0.632 \ln K+0.452 \ln L, \quad R^{2}=0.38
$$

where $\operatorname{Var}\left(\hat{\beta}_{K}\right)=0.160, \operatorname{Var}\left(\hat{\beta}_{L}\right)=0.045, \operatorname{Cov}\left(\beta_{K}, \beta_{L}\right)=0.0225$. The sample size is 100 . Test the following hypotheses at the $5 \%$ level of significance.
(a) $(7 \%) \beta_{K}=\beta_{L}$.
(b) $(8 \%)$ There are constant returns to scale.
3. $(25 \%)$ Let the multiple regression model be written as

$$
\underset{n \times 1}{y}=\underset{n \times K}{n_{K \times 1}} \underset{n \times 1}{\quad} \quad \underset{n}{u}
$$

where $y$ is the column vector of the dependent variable, $X$ is the matrix of explanatory variables, $\beta$ is the column vector of the $K$ coefficients and $u$ is the column vector of the error terms. Assume that $\mathrm{E}(u)=0$ and the covariance matrix of $u$ is $\mathrm{E}\left(u u^{\prime}\right)=\sigma^{2} I$. It is known that the OLS estimator of $\beta$ is $\hat{\beta}=\left(X^{\prime} X\right)^{-1} X^{\prime} Y$.
Let $\hat{\beta}^{*}=\hat{\beta}+C y$ be another linear estimator of $\beta$, where $C$ is a $K \times n$ matrix.
(a) $(5 \%)$ For $\hat{\beta}^{*}$ to be unbiased, what restriction do we need to impose on $C$ ?
(b) $(10 \%)$ Derive the covariance matrix of $\hat{\beta}^{*}, \operatorname{Var}\left(\hat{\beta}^{*}\right)$ ?
(c) $(10 \%)$ Compare $\operatorname{Var}\left(\hat{\beta}^{*}\right)$ with $\operatorname{Var}(\hat{\beta})$. What theorem can we have from the result of comparing the covariance matrix of the two estimators.
4. $(10 \%)$ Suppose we have the following model,

$$
Y_{i}=\beta_{1}+\beta_{2} D_{i}+\beta_{3} X_{i}+\beta_{4}\left(D_{i} * X_{i}\right)+u_{i}
$$

where $Y_{i}$ is the annual salary of a worker, $X_{i}$ is the years of working experience, and $D_{i}$ is dummy variable with

$$
\begin{aligned}
D_{i} & =1 \text { if male } \\
& =0 \text { if female }
\end{aligned}
$$

(a) (5\%) What is the difference of male and female average salaries in terms of the regression coefficient and other variables?
(b) $(5 \%)$ Suppose the dummy variable is defined as $D_{i}=1$ if female and $D_{i}=$ -1 if male, then what is the difference of male and female average salaries with 10 years of working experience.
5. $(30 \%)$ You are given a data set consists of the following 3 variables (1) schyr: one's years of schooling (2) sex : 1 for male and 0 for female (3) fschyr: father's years of schooling.
(a) (5\%) Write down the necessary Stata commands to create a new dummy variable $c$ which indicates that an individual has at least 16 years of schooling (college education).
(b) (5\%) When you regress $c$ on sex and the constant term, what is the Stata command for doing this?
(c) $(5 \%)$ you have the following result from (b),

$$
\begin{aligned}
c= & .2622419-.0345699 \text { sex } \\
& (.0073647)(.0099553)
\end{aligned}
$$

with standard errors in the parentheses. What is the proportion of male who has a college education? What is the proportion of female who has a college education?
(d) (5\%) Is man more or less likely than woman to have a college education? Is it significant at $5 \%$ level of significance?

Now you add $f$ schyr and the interaction between sex and fschyr as the explanatory variables and have the following results.


| $c$ | Coef. | Std. Err. | t | $P>\|t\|$ | [95\% Conf. Interval] |  |
| ---: | ---: | :--- | ---: | ---: | ---: | ---: |
| sex | .0030215 | .0256093 | 0.12 | 0.906 | -.0471799 | .053223 |
| fschyr | .0387226 | .001979 | 19.57 | 0.000 | .0348432 | .042602 |
| sex $*$ fschyr | -.0038253 | .0026601 | -1.44 | 0.150 | -.0090398 | .0013893 |
| cons. | -.0856444 | .0191131 | -4.48 | 0.000 | -.1231115 | -.0481774 |

(e) (5\%) Draw the regression lines of $c$ on $f s c h y r$ for male and female separately.
(f) (5\%) Is the effect of father's schooling on $c$ significantly different for son and daughter?

## Suggested Answers for Midterm Exam. II

Dec. 20, 2002
1a True. Since $A$ is symmetric, then $A^{\prime}=A .\left(A^{-1}\right)^{\prime}=\left(A^{\prime}\right)^{-1}=A^{-1}$, so $A^{-1}$ is also symmetric.

1b False. $R^{2}$ will definitely decrease but $\bar{R}^{2}$ will not necessarily decrease since $\bar{R}^{2}=$ $1-\left(1-R^{2}\right) \frac{n-1}{n-K}, \frac{\partial \bar{R}^{2}}{\partial K}=\frac{n-1}{n-K} \frac{\partial R^{2}}{\partial K}-\left(1-R^{2}\right) \frac{n-1}{(n-K)^{2}} \gtrless 0$

1c True. To minimize $\sum_{i=1}^{n} \hat{u}_{i}^{2}=\sum_{i=1}^{n}\left(Y_{i}-\hat{\beta} X_{i}\right)^{2}$, the first order condition is $\sum_{i=1}^{n}(-2)\left(Y_{i}-\hat{\beta} X_{i}\right) X_{i}=0$. Therefore, the OLS estimator of $\beta$ is $\hat{\beta}=\frac{\sum_{i=1}^{n} X_{i} Y_{i}}{\sum_{i=1}^{n} X_{i}^{2}}$.

1d False. In general $A B \neq B A$.
2a $\operatorname{Var}\left(\beta_{K}-\beta_{L}\right)=\operatorname{Var}\left(\hat{\beta}_{K}\right)+\operatorname{Var}\left(\hat{\beta}_{L}\right)-2 \operatorname{Cov}\left(\beta_{K}, \beta_{L}\right)=0.16+0.045-2 * 0.0225=$ 0.16. $t=\frac{0.632-0.452}{\sqrt{0.16}}=\frac{0.18}{0.4}=0.45$, insignificant.

2b $\operatorname{Var}\left(\beta_{K}+\beta_{L}\right)=\operatorname{Var}\left(\hat{\beta}_{K}\right)+\operatorname{Var}\left(\hat{\beta}_{L}\right)+2 \operatorname{Cov}\left(\beta_{K}, \beta_{L}\right)=0.16+0.045+2 * 0.0225=$ 0.25. $t=\frac{0.632+0.452-1}{\sqrt{0.25}}=\frac{0.084}{0.5}=0.168$, insignificant.

3a $\mathrm{E}\left[\hat{\beta}^{*}\right]=\mathrm{E}[\hat{\beta}+C X \beta+u]=\beta+C X \beta$. For $\hat{\beta}^{*}$ to be unbiased, we need $C X=0$.
$3 b$

$$
\begin{aligned}
\hat{\beta}^{*} & =\beta+\left(X^{\prime} X\right)^{-1} X^{\prime} u+C u \\
\operatorname{Var}\left(\hat{\beta}^{*}\right) & =\mathrm{E}\left[\left(\hat{\beta}^{*}-\beta\right)\left(\hat{\beta}^{*}-\beta\right)^{\prime}\right] \\
& =\mathrm{E}\left\{\left[\left(X^{\prime} X\right)^{-1} X^{\prime} u+C u\right]\left[\left(X^{\prime} X\right)^{-1} X^{\prime} u+C u\right]^{\prime}\right\} \\
& \left.=\mathrm{E}\left\{\left[\left(X^{\prime} X\right)^{-1} X^{\prime} u+C u\right]\left[u^{\prime} X X^{\prime} X\right)^{-1}+u^{\prime} C^{\prime}\right]\right\} \\
& \left.\left.=\mathrm{E}\left[\left(X^{\prime} X\right)^{-1} X^{\prime} u u^{\prime} X X^{\prime} X\right)^{-1}+C u u^{\prime} X X^{\prime} X\right)^{-1}+\left(X^{\prime} X\right)^{-1} X^{\prime} u u^{\prime} C^{\prime}+C u u^{\prime} C^{\prime}\right] \\
& =\sigma^{2}\left(X^{\prime} X\right)^{-1}+0+0+\sigma^{2} C C^{\prime} \geq \operatorname{Var}(\hat{\beta})
\end{aligned}
$$

3c $\operatorname{Var}\left(\hat{\beta}^{*}\right) \geq \operatorname{Var}(\hat{\beta})$. This is the Gauss-Markov theorem, the OLS estimator of $\beta-\hat{\beta}$ is the most efficient estimator among the class of the linear unbiased estimators.
$4 \mathbf{a}$

$$
\begin{array}{r}
\mathrm{E}\left(Y_{i} \mid D_{i}=1\right)=\beta_{1}+\beta_{2}+\beta_{3} X_{i}+\beta_{4} X_{i} \\
\mathrm{E}\left(Y_{i} \mid D_{i}=0\right)=\beta_{1}+\beta_{3} X_{i} \\
\mathrm{E}\left(Y_{i} \mid D_{1}=1\right)-\mathrm{E}\left(Y_{i} \mid D_{i}=0\right)=\beta_{2}+\beta_{4} X_{i}
\end{array}
$$

4b

$$
\begin{aligned}
\mathrm{E}\left(Y_{i} \mid D_{i}=-1, X_{i}=10\right) & =\beta_{1}-\beta_{2}+\beta_{3} * 10-\beta_{4} * 10 \\
\mathrm{E}\left(Y_{i} \mid D_{i}=1, X_{i}=10\right) & =\beta_{1}+\beta_{2}+\beta_{3} * 10+\beta_{4} * 10 \\
\mathrm{E}\left(Y_{i} \mid D_{1}=-1, X_{i}=10\right)-\mathrm{E}\left(Y_{i} \mid D_{i}\right. & \left.=1, X_{i}=10\right)=-2 \beta_{2}-20 \beta_{4}
\end{aligned}
$$

5a generate $c=1$ if schyr $>=16$
replace $c=0$ if schyr==.
5b regress c sex
5c male: $26.2-3.5=22.7 \%$, female: $26.2 \%$.
5d $t=\frac{-0.0345699}{0.0099553}=-3.47$. Man is significantly less likely to have a college education.

5e |  | son | daughter |
| :---: | :---: | :---: |
|  | intercept | -0.083 |
| slope | 0.035 | 0.039 |

## 5f

Since the coefficient of sex $* f s c h y r$ is not significant, the effect of father schooling on $c$ is not significant different for son and daughter.

