

Adaptive combination of Bayes factors method as a powerful polygenic test for gene-environment interactions when external information is unavailable

Wan-Yu Lin, Ching-Chieh Huang, Yu-Li Liu, Shih-Jen Tsai, Po-Hsiu Kuo

This work has been published by *Briefings in Bioinformatics*, with R code downloaded from <http://homepage.ntu.edu.tw/~linwy>

Institute of Epidemiology and Preventive Medicine, College of Public Health, National Taiwan University, Taipei, Taiwan

Background: The exploration of “gene-environment interactions” (GxE) is important for disease prediction and prevention. The scientific community usually uses external information to construct a genetic risk score (GRS), and then tests the interaction between this GRS and an environmental factor (E). However, external genome-wide association studies (GWAS) are not always available, especially for non-Caucasian ethnicity. Although GRS is an analysis tool to detect GxE in GWAS, its performance remains unclear when there is no external information.

Methods: Our “adaptive combination of Bayes factors method” (ADABF) can aggregate GxE signals and test the significance of GxE by a polygenic test. We here explore a powerful polygenic approach for GxE when external information is unavailable, by comparing our ADABF with the GRS based on marginal effects of SNPs (GRS-M) and GRS based on SNPxE interactions (GRS-I).

Conclusions: ADABF is the most powerful method in the absence of SNP main effects, whereas GRS-M is generally the best test when SNP main effects exist. GRS-I is the least powerful test due to its data-splitting strategy. Furthermore, we apply these methods to Taiwan Biobank data. ADABF and GRS-M identified gene-alcohol and gene-smoking interactions on blood pressure (BP). BP-increasing alleles elevate more BP in drinkers (smokers) than in nondrinkers (nonsmokers). This work provides guidance to choose a polygenic approach to detect GxE when external information is unavailable.

Methods

Adaptive combination of Bayes factors method

A pruning stage:
A screening stage:

Each additional SBP-increasing allele is associated with ~0.20 mm Hg higher SBP in drinkers than in nondrinkers.

Moreover, to improve the statistical power of $G \times E$ tests, the remained SNPs are then screened according to their marginal associations with the phenotype. The generalized linear model (GLM) for the l^{th} SNP ($l = 1, \dots, L$) is described as follows:

$$g[E(Y_i)] = \beta_0 + \beta_{G_l} G_{il} + \beta'_X X_i, i = 1, \dots, n, \quad (1)$$

where $g[\cdot]$ is the link function; Y_i is the phenotype, G_{il} is the number of minor alleles at the l^{th} SNP (0, 1 or 2) and X_i is the vector of covariates of the i^{th} subject. In this screening stage, we test $H_0: \beta_{G_l} = 0$ versus $H_1: \beta_{G_l} \neq 0$ ($l = 1, \dots, L$). The SNPs passing the screening at the desired significance level ($P < 0.05$) are then analyzed using ADABF. This screening stage that reduces the number of SNPs tested for interactions can substantially increase the power of genome-wide $G \times E$ studies

Suppose that in a GWAS there are L autosomal SNPs retained after the pruning and screening stages. We assess the interaction between the l^{th} SNP ($l = 1, \dots, L$) and E by the following GLM:

$$g[E(Y_i)] = \beta_0 + \beta_{G_l} G_{il} + \beta_E E_i + \beta_{GE_l} G_{il} E_i + \beta'_X X_i, i = 1, \dots, n; \quad (2)$$

where E_i is the environmental factor (E) of the i^{th} subject, and the other notations have been described under Equation (1). Let $\hat{\beta}_{GE_l}$ be the maximum likelihood estimate (MLE) of β_{GE_l} . According to the asymptotic normality of MLE, $\hat{\beta}_{GE_l}$ follows a normal distribution with a mean of β_{GE_l} and a variance of V_l , i.e. $\hat{\beta}_{GE_l} \sim N(\beta_{GE_l}, V_l)$.

To test whether the l^{th} SNP interacts with E , the hypothesis is $H_{0,l}: \beta_{GE_l} = 0$ versus $H_{1,l}: \beta_{GE_l} \neq 0$ ($l = 1, \dots, L$). The BF is described as follows

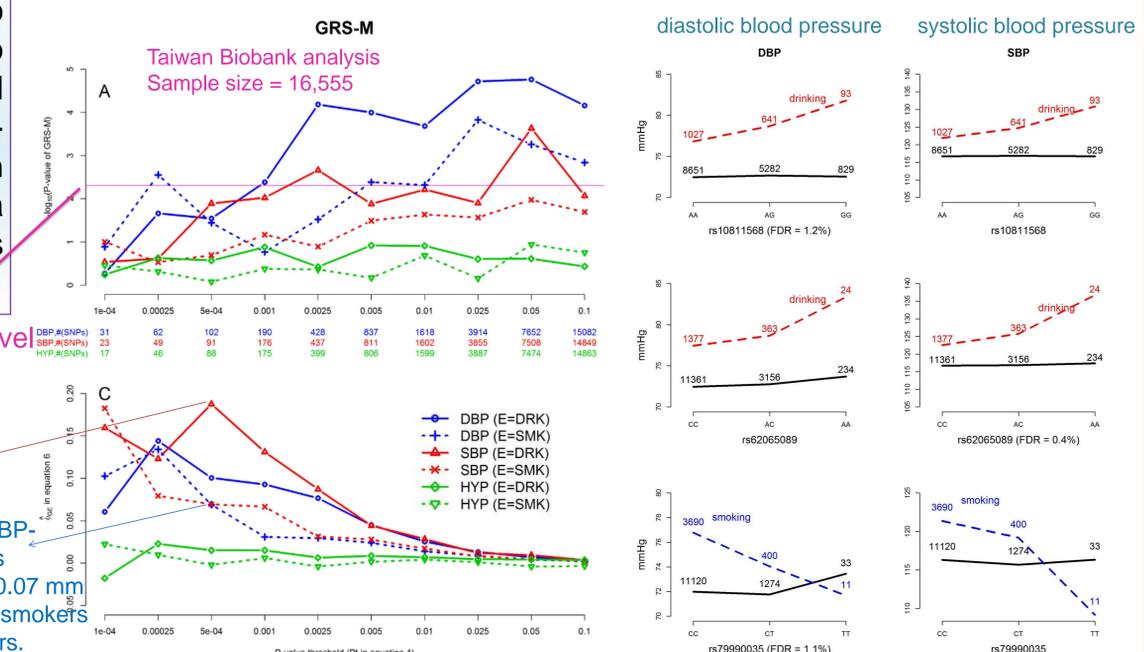
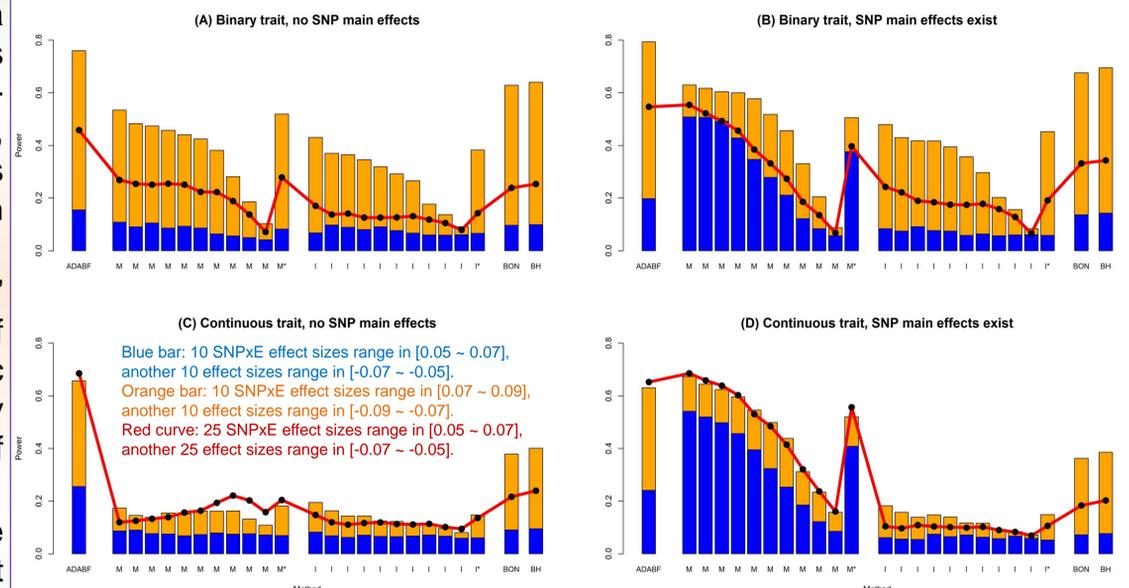
$$BF_l = \frac{\Pr(\text{Data}|H_{1,l})}{\Pr(\text{Data}|H_{0,l})} = \sqrt{\frac{\hat{V}_l}{\hat{V}_l + W}} \exp\left(\frac{\hat{\beta}_{GE_l}^2 W}{2\hat{V}_l(\hat{V}_l + W)}\right), l = 1, \dots, L, \quad (3)$$

where $\hat{\beta}_{GE_l}$ and \hat{V}_l have been estimated from the GLM in Equation (2).

$-\log_{10}(0.05/10) = 2.3$, the significance level adjusted for testing 10 times.

Each additional DBP-increasing allele is associated with ~0.07 mm Hg higher DBP in smokers than in nonsmokers.

Blue bar: 10 SNPxE odds ratios (ORs) range in [1.2 ~ 1.4], another 10 ORs range in [0.71 ~ 0.83].
Orange bar: 10 SNPxE ORs range in [1.4 ~ 1.6], another 10 ORs range in [0.63 ~ 0.71].
Red curve: 25 SNPxE ORs range in [1.2 ~ 1.4], another 25 ORs range in [0.71 ~ 0.83].



	ADABF
SNP _{alcohol} on DBP (based on 7,652 SNPs)	
P-value	< 10 ⁻⁵
SNP found to have interaction with alcohol consumption	rs10811568 (Resampling FDR = 1.2%)
SNP _{alcohol} on SBP (based on 7,508 SNPs)	
P-value	< 10 ⁻⁵
SNP found to have interaction with alcohol consumption	rs62065089 (Resampling FDR = 0.4%)
SNP _{alcohol} on HYP (based on 7,474 SNPs)	
P-value	9.8 × 10 ⁻⁴
SNP found to have interaction with alcohol consumption	---
SNP _{smoking} on DBP (based on 7,652 SNPs)	
P-value	5.9 × 10 ⁻⁴
SNP found to have interaction with smoking	rs79990035 (Resampling FDR = 1.1%)
SNP _{smoking} on SBP (based on 7,508 SNPs)	
P-value	0.1573
SNP found to have interaction with smoking	---
SNP _{smoking} on HYP (based on 7,474 SNPs)	
P-value	0.0592

GRS based on marginal effects of SNPs

We compare ADABF with GRS-M and GRS-I. Regarding GRS-M, the phenotype is first regressed on each of the L SNPs, as shown by Equation (1). The regression coefficients ($\hat{\beta}_{G_l}$ s) of the SNPs that are more associated with the phenotype (P -value less than a certain threshold) are treated as the weights of the GRS. To be specific, the pre-scaled GRS-M of the i^{th} subject is defined as follows:

$$\sum_{l=1}^L \hat{\beta}_{G_l} G_{il} I(P_{G_l} < P_t), i = 1, \dots, n; t = 1, \dots, 10, \quad (4)$$

where $\hat{\beta}_{G_l}$ is estimated by the GLM in Equation (1), G_{il} is the number of minor alleles at the l^{th} SNP of the i^{th} subject, $I(\cdot)$ is the indicator variable, P_{G_l} is the P -value of testing $H_0: \beta_{G_l} = 0$ versus $H_1: \beta_{G_l} \neq 0$ and P_t is the t^{th} P -value threshold. Most investigators use a P -value threshold to select a subset of SNPs for a GRS

We used 10 thresholds to explore the strength of GRS: 0.0001, 0.00025, 0.0005, 0.001, 0.0025, 0.005, 0.01, 0.025, 0.05 and 0.1.

$GRS_{Mi,t}^{pre}$ is then rescaled to calibrate the number of phenotype-increasing alleles

$$GRS_{Mi,t} = \frac{GRS_{Mi,t}^{pre} \times \text{number of available SNPs}}{\text{sum of } |\hat{\beta}_{G_l}| \text{ of available SNPs}} \quad (5)$$

Given the t^{th} P -value threshold ($t = 1, \dots, 10$), we calculate $GRS_{Mi,t}$ for all the n subjects, fit the following GLM, and test $H_0: \phi_{GE} = 0$ versus $H_1: \phi_{GE} \neq 0$:

$$g[E(Y_i)] = \phi_0 + \phi_C GRS_{Mi,t} + \phi_E E_i + \phi_{GE} GRS_{Mi,t} \cdot E_i + \phi'_X X_i, i = 1, \dots, n. \quad (6)$$

Because we consider 10 P -value thresholds, 10 GLMs are fitted and $H_0: \phi_{GE} = 0$ is tested 10 times.

References:

- Lin, W. Y., et al. (2018). *Briefings in Bioinformatics*, in press.
- Lin, W. Y., et al. (2017). *Scientific Reports*, 7: 13858.
- Hüls A, et al. *BMC Genetics* 2017;18: 115.
- Hüls A, et al. *BMC Genetics* 2017;18: 55.

gene-alcohol interaction > gene-smoking interaction for blood pressure levels