

Experimental Economics I

Jury Voting

Instructor: Sun-Tak Kim

National Taiwan University

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Jury Voting Model

- ▶ Three jurors $N = \{1, 2, 3\}$ responsible for deciding whether to convict or acquit a defendant.
- ▶ Collectively they choose an outcome $x \in \{c, a\}$.
- ▶ The jurors simultaneously cast ballots $v_i \in S_i = \{c, a\}$.
- ▶ The outcome is chosen by majority rule.
- ▶ Each juror is uncertain whether or not the defendant is guilty (G) or innocent (I).
- ▶ So the set of state variables is $\Omega = \{G, I\}$.
- ▶ Each juror assigns prior prob. $\pi > 1/2$ to state G.
- ▶ If the defendant is guilty, the jurors receive 1 unit of utility from convicting and 0 from acquitting; if the defendant is innocent, the jurors receive 1 unit from acquitting and 0 from convicting;

$$\left. \begin{aligned} u(c|G) &= u(a|I) = 1 \\ u(a|G) &= u(c|I) = 0 \end{aligned} \right\}$$

Jury Voting Model

- ▶ Absent any additional information, each juror receives an expected utility of π from a guilty verdict and $1 - \pi$ from an acquittal.
- ▶ Because $\pi > 1/2$, the Nash eq'u'm that survives the elimination of weakly dominated strategies is the one where each juror votes guilty.
- ▶ Now, before voting, each juror receives a **private signal** about the defendant's guilt $\theta_i \in \{0, 1\}$.
- ▶ The signal is **informative** so that a juror is more likely to receive the signal $\theta_i = 1$ when the defendant is guilty than when the defendant is innocent.
- ▶ Assume the prob. of receiving a “guilty” signal ($\theta_i = 1$) when the defendant is guilty is the same as that of receiving an “innocent” signal ($\theta_i = 0$) when the defendant is innocent. *Symmetry*
- ▶ Formally, let $\Pr(\theta_i = 1|\omega = G) = \Pr(\theta_i = 0|\omega = I) = p > 1/2$ so that $\Pr(\theta_i = 0|\omega = G) = \Pr(\theta_i = 1|\omega = I) = 1 - p$.
- ▶ Conditional on a state, each signal for an individual is independent with each other (signals are “**conditionally independent**”).

Sincere Voting Strategy

- ▶ After receiving her signal, voter i selects her **vote** $v_i(\theta_i)$ to maximize the prob. of a correct decision - conviction of the guilty and acquittal of the innocent.
- ▶ Suppose that each voter uses the **sincere strategy** $v_i(1) = c$ and $v_i(0) = a$.
- ▶ The sincere strategy calls for a vote to convict upon receipt of a guilty signal and a vote to acquit upon an innocent signal.
- ▶ **Sincere** strategies constitute a Bayesian Nash equ'm (**BNE**) only if voter 1 is willing to use this strategy when she believes that voters 2 and 3 also use it.
- ▶ Given these conjectures, the expected utility (**EU**) **of voting to convict** is

$$\begin{aligned}
 & \text{One other vote } c \\
 & \Pr(\theta_2 = 1, \theta_3 = 0; \omega = G | \theta_1) + \Pr(\theta_2 = 0, \theta_3 = 1; \omega = G | \theta_1) \\
 + & \Pr(\theta_2 = 1, \theta_3 = 1; \omega = G | \theta_1) + \Pr(\theta_2 = 0, \theta_3 = 0; \omega = I | \theta_1). \\
 & \text{Both vote } c \qquad \qquad \qquad \text{None vote } c
 \end{aligned}$$

Sincere Voting Strategy

- Suppose that juror 1 receives $\theta_1 = 1$.
- In this case, Bayes' rule yields

$$\Pr(\theta_2 = 1, \theta_3 = 0; \omega = G | \theta_1 = 1) \stackrel{\text{event B}}{=} \frac{\Pr(A|B) //}{\Pr(B) //}$$

$$\stackrel{\text{event A}}{=} \frac{\Pr(\theta_2 = 0, \theta_3 = 1; \omega = G | \theta_1 = 1)}{\pi p + (1 - \pi)(1 - p)}$$

$$\frac{\Pr(\theta_1=1, \theta_2=1, \theta_3=0 | G) \Pr(G)}{\Pr(\theta_1=1 | G) \Pr(G) + \Pr(\theta_1=1 | I) \Pr(I)}$$

and

$$\Pr(\theta_2 = 1, \theta_3 = 0; \omega = I | \theta_1 = 1)$$

$$= \Pr(\theta_2 = 0, \theta_3 = 1; \omega = I | \theta_1 = 1) = \frac{(1 - \pi)p(1 - p)^2}{\pi p + (1 - \pi)(1 - p)}$$

- Thus, $v_i(1) = c$ is optimal for juror 1 if

$$2 \frac{\pi p^2(1 - p)}{\pi p + (1 - \pi)(1 - p)} \geq 2 \frac{(1 - \pi)p(1 - p)^2}{\pi p + (1 - \pi)(1 - p)}$$

$$\Rightarrow 2\pi p^2(1 - p) \geq (1 - \pi)p(1 - p)^2 + \pi p^2(1 - p)$$

Sincere Voting Strategy

- ▶ After simplifying and rearranging, this inequality becomes

$$P_r(G | \theta_i=1, \theta_j=1, \theta_k=0) \stackrel{\text{(I am pivotal!)}}{=} \frac{\pi p^2(1-p)}{\pi p^2(1-p) + (1-\pi)p(1-p)^2} \geq \frac{1}{2}.$$

- ▶ LHS is just the **conditional prob. of guilt** given two signals of $\theta = 1$ and one signal of $\theta = 0$.
- ▶ In other words, agent 1 wants to **vote to convict** if she believes that the defendant is more likely to be guilty than innocent, conditional on her signal and the belief that she is pivotal.
- ▶ Similarly, the requirement for a **vote of innocence** conditional on a signal of 0 is

$$\frac{\pi p(1-p)^2}{\pi p(1-p)^2 + (1-\pi)p^2(1-p)} \leq \frac{1}{2}.$$

- ▶ To sum, in any BNE in which voting corresponds to the private signals,
 1. Conditional on the supposition that i is pivotal and observes $\theta_i = 1$, the **posterior prob. of guilt** is greater than **1/2**; and
 2. Conditional on the supposition that i is pivotal and observes $\theta_i = 0$, the **posterior prob. of guilt** is less than **1/2**.

Asymmetric Signal

- ▶ Thus, if sincere voting is incentive compatible, then

$$\frac{1-p}{p} \leq \frac{\pi}{1-\pi} \leq \frac{p}{1-p}.$$

- ▶ E.g., if $\pi > p$, then sincere voting is not incentive compatible.
- ▶ Under majority rule and symmetric signal precision (and equal prior $\pi = 1/2$), sincere voting obtains in equilibrium (if $p > 1/2$).
- ▶ Alternative way to obtain an *insincere/strategic voting equilibrium* is to introduce asymmetric signal:

$$\begin{aligned} p &\equiv \Pr(\theta_i = 1 | \omega = G), & q &\equiv \Pr(\theta_i = 0 | \omega = I), \\ 1-p &= \Pr(\theta_i = 0 | \omega = G), & 1-q &= \Pr(\theta_i = 1 | \omega = I), \end{aligned}$$

and we have here $1 > p > q > 1/2$.

- ▶ Then, the posterior probabilities (with equal prior $\pi = 1/2$) are

$$\Pr[\omega = G | \theta_i = 1] = \frac{p}{p + (1-q)}, \quad \Pr[\omega = I | \theta_i = 0] = \frac{q}{(1-p) + q}.$$

Strategic Voting Equ'm

- ▶ Define $\sigma(s) \equiv$ prob. of voting one's signal, $s = 0, 1$.
- ▶ Typically, we have in equ'm; $\sigma(1) \in (0, 1)$ and $\sigma(0) = 1$. (semi-pooling equily)
- ▶ Then,

$$\Pr[c|\omega = G] = p\sigma(1) + (1-p)(1-\sigma(0)) = p\sigma(1),$$

$$\Pr[a|\omega = G] = p(1-\sigma(1)) + (1-p)\sigma(0) = p(1-\sigma(1)) + (1-p),$$

$$\Pr[c|\omega = I] = (1-q)\sigma(1) + q(1-\sigma(0)) = (1-q)\sigma(1),$$

$$\Pr[a|\omega = I] = (1-q)(1-\sigma(1)) + q\sigma(0) = (1-q)(1-\sigma(1)) + q,$$

- ▶ Since the equ'm strategy requires randomization upon signal $s = 1$,

Indifference $\Rightarrow \Pr[\omega = G|\theta_i = 1] \Pr[\text{Piv}|\omega = G] - \Pr[\omega = I|\theta_i = 1] \Pr[\text{Piv}|\omega = I] = 0,$

where $\Pr[\text{Piv}|\omega]$ is the prob. a vote is pivotal at state ω :

$$\begin{aligned} \Pr[\text{Piv}|\omega = G] &= \binom{2}{1} \Pr[c|\omega = G] \Pr[a|\omega = G] \\ &= [p\sigma(1)][p(1-\sigma(1)) + (1-p)], \end{aligned}$$

(One vote c, the other vote a)

Strategic Voting Equ'm

$$\begin{aligned}\Pr[\text{Piv}|\omega = I] &= \binom{2}{1} \Pr[c|\omega = I] \Pr[a|\omega = I] \\ &= [(1-q)\sigma(1)][(1-q)(1-\sigma(1)) + q]\end{aligned}$$

- ▶ Thus we solve for $\sigma(1)$ from the above equation.
- ▶ Since $\sigma(0) = 1$, we finally check whether

$$\Pr[\omega = I|\theta_i = 0] \Pr[\text{Piv}|\omega = I] - \Pr[\omega = G|\theta_i = 0] \Pr[\text{Piv}|\omega = G] > 0$$

when $\Pr[\text{Piv}|\omega]$ is evaluated at $\sigma(1)$ that solves the indifference condition.

- ▶ For example, when $p = 0.9$ and $q = 0.6$, $\sigma(1) = 0.9774$
- ▶ Under fixed (p, q) , $\sigma(1)$ typically **decreases** as n gets larger.

- ▶ Austen-Smith & Banks (1996) show that in many cases the **sincere strategy is inconsistent** with equilibrium behavior.
- ▶ It is easy to find parameters π and p for which one of the necessary conditions does not hold.
- ▶ There are alternative strategies jurors might choose.
- ▶ Jurors can randomize for some signals, vote the same way regardless of their signal, or use different strategies than other jurors use.
- ▶ Feddersen & Pesendorfer (1998) consider the properties of equ'a of this game when one varies the voting rule and number of jurors.

Jury Voting with a Continuum of Signals

- ▶ Instead of receiving a binary signal, each juror now receives a signal $\theta_i \in [0, 1]$ where θ_i is drawn from a **conditional distribution** $F(\theta_i|\omega)$.
- ▶ This distribution function is associated with a different density function $f(\theta_i|\omega)$ that satisfies the **monotone likelihood ratio condition**.
- ▶ A conditional density function satisfies the **strict monotone likelihood ratio condition (SMLR)** if $\frac{f(\theta_i|G)}{f(\theta_i|I)}$ is a **strictly monotone** function of θ_i on $[0, 1]$.
- ▶ To see why this assumption is important, note that Bayes' rule implies that

$$\begin{aligned}\Pr(G|\theta_i) &= \frac{f(\theta_i|G)\pi}{f(\theta_i|G)\pi + f(\theta_i|I)(1 - \pi)} \\ &= \frac{\frac{f(\theta_i|G)}{f(\theta_i|I)}\pi}{\frac{f(\theta_i|G)}{f(\theta_i|I)}\pi + (1 - \pi)}.\end{aligned}$$

- ▶ Accordingly, $\Pr(G|\theta_i)$ is **increasing in θ_i** if & only if **$f(\theta_i|G)/f(\theta_i|I)$ is increasing in θ_i** .
- ▶ Thus, the SMLR condition implies that higher signals correspond to higher posterior probabilities that $\omega = G$.

Jury Voting with a Continuum of Signals

- ▶ To keep matters simple, we focus exclusively on **symmetric strategies** where voters who receive the same signal choose the same strategy.
- ▶ A symmetric strategy profile is, therefore, a mapping $v_i(\theta_i) : [0, 1] \rightarrow \{c, a\}$.
- ▶ As in the binary signal case, BNE strategies are those that are optimal when each agent acts conditionally on her private information and the conjecture that she is pivotal.
- ▶ An agent votes to convict if she thinks the **prob. of guilt** is no less than $1/2$ and she votes to acquit if she thinks the **prob. of guilt** is no more than $1/2$.
- ▶ Because higher signals are better indicators of guilt, a natural conjecture is that the strategy must be **weakly increasing**.
- ▶ For low values of θ_i an acquittal vote is cast and for high values of θ_i a conviction vote is cast.

Cut Point Strategy

- ▶ A monotone strategy of this form can be characterized by a **cut point** $\hat{\theta} \in [0, 1]$.
- ▶ Assume that agents $i \in N \setminus i$ use the monotone strategy

$$v_i(\theta_i) = \begin{cases} c & \text{if } \theta_i \geq \hat{\theta} \\ a & \text{if } \theta_i < \hat{\theta} \end{cases}$$

- ▶ If all players other than i use this cut point strategy, the **posterior prob. of $\{\omega = G\}$ given signal θ_i and the event that i is pivotal** is given by

$$\begin{aligned} & \Pr(G|piv, \theta_i; \hat{\theta}) \\ &= \frac{\pi f(\theta_i|G)F(\hat{\theta}|G)^{N-r}[1 - F(\hat{\theta}|G)]^{r-1}}{\pi f(\theta_i|G)F(\hat{\theta}|G)^{N-r}[1 - F(\hat{\theta}|G)]^{r-1} + (1 - \pi)f(\theta_i|I)F(\hat{\theta}|I)^{N-r}[1 - F(\hat{\theta}|I)]^{r-1}} \end{aligned}$$

- ▶ This prob. is a function of the parameter $\hat{\theta}$.
- * Here we assume **r-rule**, so we require **r or more** votes for conviction (**majority** rule if $r = (N + 1)/2$ and **unanimity** rule if $r = N$).

Cut Point Equilibrium

- In this model the existence of a symmetric equ'm in which voters use a cut point hinges on finding a value of $\hat{\theta}$ s.t.

$$\Pr(G|piv, \hat{\theta}; \hat{\theta}) = \frac{1}{2}$$

and demonstrating that $\Pr(G|piv, \theta_i; \hat{\theta}) \leq \frac{1}{2}$ if $\theta_i < \hat{\theta}$ and $\Pr(G|piv, \theta_i; \hat{\theta}) \geq \frac{1}{2}$ if $\theta_i > \hat{\theta}$.

- Although analysis of examples is cumbersome, it is easy to derive conditions on the primitives of the game to ensure that such a $\hat{\theta} \in (0, 1)$ exists.
- First, $\Pr(G|piv, \theta_i; \hat{\theta}) \geq \frac{1}{2}$ if & only if

$$\begin{aligned}
 H(\theta_i) &= \frac{\pi f(\theta_i|G) F(\hat{\theta}|G)^{N-r} [1 - F(\hat{\theta}|G)]^{r-1}}{(1 - \pi) f(\theta_i|I) F(\hat{\theta}|I)^{N-r} [1 - F(\hat{\theta}|I)]^{r-1}} \\
 &= \frac{f(\theta_i|G)}{f(\theta_i|I)} \frac{\pi F(\hat{\theta}|G)^{N-r} [1 - F(\hat{\theta}|G)]^{r-1}}{(1 - \pi) F(\hat{\theta}|I)^{N-r} [1 - F(\hat{\theta}|I)]^{r-1}}
 \end{aligned}$$

$H(\theta) < 1$
 $\Rightarrow H(\theta) > 1$
 $\Rightarrow \exists \hat{\theta} \in (0, 1)$
 s.t. $H(\hat{\theta}) = 1$

Existence of Cut Point Equilibrium

- ▶ SMLR then implies that if $\Pr(G|piv, \hat{\theta}_i; \hat{\theta}) = 1/2$ then $\theta_i < \hat{\theta}$ implies $\Pr(G|piv, \theta_i; \hat{\theta}) \leq 1/2$ and $\theta_i > \hat{\theta}$ implies $\Pr(G|piv, \theta_i; \hat{\theta}) \geq 1/2$.
- ▶ If $\Pr(G|piv, 0; \mathbf{0}) \leq 1/2 \leq \Pr(G|piv, 1; \mathbf{1})$ then the intermediate value theorem implies that such a cut point exists b/c the function $\Pr(G|piv, \cdot; \cdot)$ is continuous.
- ▶ For a large class of games these boundary conditions are satisfied.
- ▶ In the simple binary signal model, equ'a where everyone uses the same rule and voting is determined by private information may not exist.
- ▶ This type of equ'm generally exists in the continuum model, however.
- ▶ Using the **binary** model, **Feddersen & Pesendorfer (1998)** show that the **unanimity** rule is a uniquely bad way to aggregate information for large populations b/c in equ'm voters condition on the assumption that everyone else is voting to convict.
- ▶ In the **continuum** model, **Meirowitz (2002)** shows that the unanimity rule often turns out to be as good as the other voting rules.

Voluntary Voting Model

- ▶ Two **candidates**, **A** and **B**, in **majority** voting election.
- ▶ Two equally likely **states** of nature, **α** and **β** .
- ▶ **A** is the **better choice** in state **α** and **B**, in state **β** .
 - In state α , payoff is **1** if A is elected and **0** if B is elected; vice versa in state β .

- ▶ The **size of the electorate** is a random variable, distributed according to a **Poisson distribution** with mean **n** .

The Poisson distribution is for analytic convenience.

- The probability that there are exactly **m** voters is **$e^{-n} n^m / m!$** .
- ▶ Prior to voting, each voter receives a **private signal** **S_i** ; regarding the true state of nature, either **a** or **b** ; **$\Pr[a|\alpha] = r$** and **$\Pr[b|\beta] = s$** ; the **posteriors** given by

$$q(\alpha|a) = \frac{r}{r + (1 - s)}, \quad q(\beta|b) = \frac{s}{s + (1 - r)}.$$

- **$r \geq s > 1/2$** implies **$q(\alpha|a) \leq q(\beta|b)$** .

Pivotal Events

- ▶ **Event** (j, k) , j votes for **A** and k votes for **B**.
- ▶ An event is **pivotal** for A if a single additional vote for A changes the outcome, written Piv_A . *The additional vote makes or breaks a tie.*
- ▶ Under majority rule, one additional vote for A makes a difference only if **(i)** there is a **tie**; or **(ii)** A has **one vote less** than B.

$$T = \{(k, k) : k \geq 0\}, \quad T_{-1} = \{(k-1, k) : k \geq 1\}, \quad Piv_A = T \cup T_{-1}$$

- ▶ Similarly, $Piv_B = T \cup T_{+1}$, $T_{+1} = \{(k, k-1) : k \geq 1\}$.
- ▶ σ_A, σ_B are the **expected number of votes for A, B** in state α ; τ_A, τ_B are the **expected number of votes for A, B** in state β .
- ▶ With **abstention allowed**, $\sigma_A + \sigma_B \leq n$, $\tau_A + \tau_B \leq n$ (equality w/o abstention).

Pivotal Events

- ▶ If the realized electorate is of size m with k votes for **A** and l votes for **B** ($m - k - l$ abstention),

$$\Pr[(k, l)|\alpha] = e^{-\sigma_A} \frac{\sigma_A^k}{k!} e^{-\sigma_B} \frac{\sigma_B^l}{l!}.$$

- * For the probability of the event (k, l) in state β , replace σ by τ .

$$\Pr[T|\alpha] = e^{-\sigma_A - \sigma_B} \sum_{k=0}^{\infty} \frac{\sigma_A^k}{k!} \frac{\sigma_B^k}{k!},$$

$$\Pr[T_{-1}|\alpha] = e^{-\sigma_A - \sigma_B} \sum_{k=1}^{\infty} \frac{\sigma_A^{k-1}}{(k-1)!} \frac{\sigma_B^k}{k!},$$

$$\Pr[\text{Piv}_A|\alpha] = \frac{1}{2} \Pr[T|\alpha] + \frac{1}{2} \Pr[T_{-1}|\alpha] \quad \text{actually from utility } (0 \rightarrow \frac{1}{2} \text{ or } \frac{1}{2} \rightarrow 1)$$

where $\text{Piv}_A \equiv T \cup T_{-1}$ is the set of events where one additional vote for A is decisive, and we have the coefficient $1/2$ because the additional vote for A breaks a tie or leads to a tie.

Pivotal Events

- ▶ Similarly,

$$\Pr[\text{Piv}_B|\beta] = \frac{1}{2} \Pr[T|\beta] + \frac{1}{2} \Pr[T_{+1}|\beta]$$

where $\text{Piv}_B = T \cup T_{+1}$ is the set of events where one additional vote for B is decisive.

- ▶ Define *modified Bessel functions*

$$I_0(z) = \sum_{k=0}^{\infty} \frac{(z/2)^k}{k!} \frac{(z/2)^k}{k!}, \quad I_1(z) = \sum_{k=1}^{\infty} \frac{(z/2)^{k-1}}{(k-1)!} \frac{(z/2)^k}{k!}$$

and rewrite the probabilities of close elections in terms of these functions

$$\begin{aligned} \Pr[T|\alpha] &= e^{-\sigma_A - \sigma_B} I_0(2\sqrt{\sigma_A \sigma_B}) \\ \Pr[T_{\pm 1}|\alpha] &= e^{-\sigma_A - \sigma_B} \left(\frac{\sigma_A}{\sigma_B}\right)^{\pm 1/2} I_1(2\sqrt{\sigma_A \sigma_B}). \end{aligned}$$

- ▶ For z large, we also have

$$I_0(z) \approx \frac{e^z}{\sqrt{2\pi z}} \approx I_1(z).$$

Compulsory Voting

- ▶ By **compulsory voting** each voter must cast a vote for either A or B.
- ▶ Vote **sincerely** in compulsory voting equilibrium?
- ▶ Given sincere and compulsory voting, $\sigma_A = nr$, $\sigma_B = n(1-r)$, $\tau_A = n(1-s)$, $\tau_B = ns$.
- ▶ As n increases, both $\sigma \rightarrow \infty$, $\tau \rightarrow \infty$, and so the previous approximations for $I_0(z)$, $I_1(z)$ imply

$$\frac{\Pr[\text{Piv}_A|\alpha] + \Pr[\text{Piv}_B|\alpha]}{\Pr[\text{Piv}_A|\beta] + \Pr[\text{Piv}_B|\beta]} \approx \frac{e^{2n\sqrt{r(1-r)}}}{e^{2n\sqrt{s(1-s)}}} \times K(r,s) \xrightarrow{\text{(as } n \rightarrow \infty)} 0$$

where $K(r,s)$ is positive and independent of n .

- ▶ $r > s > 1/2$ also implies $s(1-s) > r(1-r)$ and so RHS goes to **zero** as n increases.

Compulsory Voting

- ▶ This implies that, when n is large and a voter is **pivotal**, state β is **infinitely more likely** than state α .
- ▶ Thus, voters with a signals will not wish to vote sincerely.

Proposition 1: *Suppose $r > s$. If voting is **compulsory**, **sincere voting is not an equilibrium** in large elections.*

- ▶ This result also holds for a fixed number of voters (Feddersen & Pesendorfer APSR 1998).

Voluntary Voting

- ▶ **Costly voting**: one's cost of voting is private info and an **independent** draw from a continuous distribution F with support $[0, 1]$ - F admits a **density $f > 0$** on $[0, 1]$.
- ▶ Voting costs are independent of the signals.
- ▶ There exists an **equilibrium** of this voluntary (and costly) voting game with the following features;
 - (i) There exists a pair of positive **threshold costs** c_a, c_b s.t. a voter with cost c and signal $i = a, b$ votes (does not abstain) if & only if $c \leq c_i$. The threshold costs determine differential **participation rates** $F(c_a) = p_a, F(c_b) = p_b$.
 - (ii) All those who vote do so **sincerely** - i.e., all those with signal a vote for **A** and those with signal b vote for **B**.

Equ'm Participation Rates

- ▶ We show that when all those who vote do so sincerely, there is an **equ'm in cutoff strategies**.
- ▶ There exists a **threshold cost** $c_a > 0$ ($c_b > 0$) s.t. all voters with signal i and cost $c \leq c_a$ ($c \leq c_b$) go to the polls and vote for **A (B)**.
→ strictly positive!
- ▶ These then determine **participation probabilities** $p_a = F(c_a)$, $p_b = F(c_b)$ for voters with signal a, b , respectively.
- ▶ Now the **expected numbers of votes** for **A, B** in state α are $\sigma_A = nrp_a$, $\sigma_B = n(1-r)p_b$; and those in state β are $\tau_A = n(1-s)p_a$, $\tau_B = nsp_b$, respectively.
- ▶ We look for participation rates p_a, p_b s.t. a voter with signal a and cost $c_a = F^{-1}(p_a)$ is indifferent b/w going to the polls and staying home;

$$(IRa) \quad U_a(p_a, p_b) \equiv q(\alpha|a) \Pr[\text{Piv}_A|\alpha] - q(\beta|a) \Pr[\text{Piv}_A|\beta] = F^{-1}(p_a)$$

benefit of voting = *cost of voting*

Equ'm Participation Rates

where the pivot probabilities are determined using the expected vote totals σ , τ .

- ▶ Similarly, a voter with signal b and cost $c_b = F^{-1}(p_b)$ must also be indifferent;

$$(IRb) \quad U_b(p_a, p_b) \equiv q(\beta|b) \Pr[Piv_B|\beta] - q(\alpha|b) \Pr[Piv_B|\alpha] = F^{-1}(p_b).$$

Proposition 2: *There exist participation rates $p_a^* \in (0, 1)$ and $p_b^* \in (0, 1)$ that simultaneously satisfy (IRa) and (IRb).*

- ▶ **Intuition for positive participation rates:** assume $p_a = 0$.
- ▶ Then the only pivotal events are $(0, 0)$ and $(0, 1)$.

Equ'm Participation Rates

- ▶ Hence conditional on being pivotal

$$\frac{\Pr[\text{Piv}_A|\alpha]}{\Pr[\text{Piv}_A|\beta]} = \frac{e^{-n(1-r)p_b}}{e^{-nsp_b}} \times \frac{1 + n(1-r)p_b}{1 + nsp_b}.$$

- ▶ The ratio of the exponential terms favors state α while the ratio of the linear terms favors state β ; and the exponential terms always dominate.
- ▶ Since state α is perceived more likely than β by a voter with signal a who is pivotal, the payoff from voting is positive. \Rightarrow Contradicting $p_a = 0$ (\rightarrow ✗)
- ▶ We also have

Lemma 1: If $r > s$, then any solution to (IRa) and (IRb) satisfies $p_a^* < p_b^*$, with equality if $r = s$.

Sincere Voting

- ▶ Given the (equ'm) participation rates, we can show that it is a best-response for every voter to vote sincerely.

- ▶ We begin with a lemma;

Lemma 2: If voting behavior is s.t. $\sigma_A > \tau_A$ and $\sigma_B < \tau_B$, then

$$\frac{\Pr[\text{Piv}_B|\alpha]}{\Pr[\text{Piv}_B|\beta]} > \frac{\Pr[\text{Piv}_A|\alpha]}{\Pr[\text{Piv}_A|\beta]} \quad (\text{Need Poisson distribution !!})$$

- ▶ On the set of “marginal” events where the vote totals are close (i.e., a voter is pivotal), **A** is more likely to be **leading** in state α and more likely to be **trailing** in state β .
- ▶ Let (p_a^*, p_b^*) be **equ'm participation** rates.
- ▶ A voter with signal a and cost $c_a^* = F^{-1}(p_a^*)$ is just indifferent b/w voting and staying home;

$$(IRa) \quad q(\alpha|a) \Pr[\text{Piv}_A|\alpha] - q(\beta|a) \Pr[\text{Piv}_A|\beta] = F^{-1}(p_a^*) > 0$$

Sincere Voting

- ▶ To show: sincere voting is optimal for a voter with signal a if others are voting sincerely;

$$(IC_a) \quad q(\alpha|a) \Pr[\text{Piv}_A|\alpha] - q(\beta|a) \Pr[\text{Piv}_A|\beta] \\ \geq q(\beta|a) \Pr[\text{Piv}_B|\beta] - q(\alpha|a) \Pr[\text{Piv}_B|\alpha].$$

- ▶ LHS is the payoff from voting for A whereas RHS is the payoff to voting for B.
- ▶ $p_a^* > 0$ combined with (IR_a) implies

$$\frac{\Pr[\text{Piv}_A|\alpha]}{\Pr[\text{Piv}_A|\beta]} > \frac{q(\beta|a)}{q(\alpha|a)}.$$

- ▶ Then by Lemma 2,

$$\frac{\Pr[\text{Piv}_B|\alpha]}{\Pr[\text{Piv}_B|\beta]} > \frac{q(\beta|a)}{q(\alpha|a)}.$$

- ▶ But then, the last inequality is equivalent to

$$q(\beta|a) \Pr[\text{Piv}_B|\beta] - q(\alpha|a) \Pr[\text{Piv}_B|\alpha] < 0. \quad (\text{---})$$

- ▶ Similarly, we combine $p_b^* > 0$, Lemma 2, and

$$(IRb) \quad q(\beta|b) \Pr[Piv_B|\beta] - q(\alpha|b) \Pr[Piv_B|\alpha] = F^{-1}(p_b^*)$$

to show

$$(ICb) \quad \begin{aligned} & q(\beta|b) \Pr[Piv_B|\beta] - q(\alpha|b) \Pr[Piv_B|\alpha] \\ & \geq q(\alpha|b) \Pr[Piv_A|\alpha] - q(\beta|b) \Pr[Piv_A|\beta]. \end{aligned}$$

Proposition 3: Under voluntary participation, *sincere voting is incentive compatible.*

- ▶ We can also show that *all equ'a* involve sincere voting (Krishna & Morgan JET 2012).