

# Experimental Economics I

## Jury Voting

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# Jury Voting Model

- ▶ Three jurors  $N = \{1, 2, 3\}$  responsible for deciding whether to convict or acquit a defendant.
- ▶ Collectively they choose an outcome  $x \in \{c, a\}$ .
- ▶ The jurors simultaneously cast ballots  $v_i \in S_i = \{c, a\}$ .
- ▶ The outcome is chosen by majority rule.
- ▶ Each juror is uncertain whether or not the defendant is guilty (G) or innocent (I).
- ▶ So the set of state variables is  $\Omega = \{G, I\}$ .
- ▶ Each juror assigns prior prob.  $\pi > 1/2$  to state G.
- ▶ If the defendant is guilty, the jurors receive 1 unit of utility from convicting and 0 from acquitting; if the defendant is innocent, the jurors receive 1 unit from acquitting and 0 from convicting;

$$\left. \begin{aligned} u(c|G) &= u(a|I) = 1 \\ u(a|G) &= u(c|I) = 0 \end{aligned} \right\}$$

# Jury Voting Model

- ▶ Absent any additional information, each juror receives an expected utility of  $\pi$  from a guilty verdict and  $1 - \pi$  from an acquittal.
- ▶ Because  $\pi > 1/2$ , the Nash eq'm that survives the elimination of weakly dominated strategies is the one where each juror votes guilty.
- ▶ Now, before voting, each juror receives a **private signal** about the defendant's guilt  $\theta_i \in \{0, 1\}$ .
- ▶ The signal is **informative** so that a juror is more likely to receive the signal  $\theta_i = 1$  when the defendant is guilty than when the defendant is innocent.
- ▶ Assume the prob. of receiving a “guilty” signal ( $\theta_i = 1$ ) when the defendant is guilty is the same as that of receiving an “innocent” signal ( $\theta_i = 0$ ) when the defendant is innocent. *Symmetry*
- ▶ Formally, let  $\Pr(\theta_i = 1|\omega = G) = \Pr(\theta_i = 0|\omega = I) = p > 1/2$  so that  $\Pr(\theta_i = 0|\omega = G) = \Pr(\theta_i = 1|\omega = I) = 1 - p$ .
- ▶ Conditional on a state, each signal for an individual is independent with each other (signals are “**conditionally independent**”).

# Sincere Voting Strategy

- ▶ After receiving her signal, voter  $i$  selects her **vote**  $v_i(\theta_i)$  to maximize the prob. of a correct decision - conviction of the guilty and acquittal of the innocent.
- ▶ Suppose that each voter uses the **sincere strategy**  $v_i(1) = c$  and  $v_i(0) = a$ .
- ▶ The sincere strategy calls for a vote to convict upon receipt of a guilty signal and a vote to acquit upon an innocent signal.
- ▶ **Sincere** strategies constitute a Bayesian Nash equ'm (**BNE**) only if voter 1 is willing to use this strategy when she believes that voters 2 and 3 also use it.
- ▶ Given these conjectures, the expected utility (**EU**) **of voting to convict** is

$$\begin{aligned}
 & \text{One other vote } c \\
 & \Pr(\theta_2 = 1, \theta_3 = 0; \omega = G | \theta_1) + \Pr(\theta_2 = 0, \theta_3 = 1; \omega = G | \theta_1) \\
 + & \Pr(\theta_2 = 1, \theta_3 = 1; \omega = G | \theta_1) + \Pr(\theta_2 = 0, \theta_3 = 0; \omega = I | \theta_1). \\
 & \text{Both vote } c \qquad \qquad \qquad \text{None vote } c
 \end{aligned}$$



# Sincere Voting Strategy

- Suppose that juror 1 receives  $\theta_1 = 1$ .
- In this case, Bayes' rule yields

$$\Pr(\theta_2 = 1, \theta_3 = 0; \omega = G | \theta_1 = 1) \stackrel{\text{event B}}{=} \frac{\Pr(A|B) //}{\Pr(B) //}$$

$$\stackrel{\text{event A}}{=} \frac{\Pr(\theta_2 = 0, \theta_3 = 1; \omega = G | \theta_1 = 1)}{\pi p + (1 - \pi)(1 - p)}$$

$$\frac{\Pr(\theta_1=1, \theta_2=1, \theta_3=0 | G) \Pr(G)}{\Pr(\theta_1=1 | G) \Pr(G) + \Pr(\theta_1=1 | I) \Pr(I)}$$

and

$$\Pr(\theta_2 = 1, \theta_3 = 0; \omega = I | \theta_1 = 1)$$

$$= \Pr(\theta_2 = 0, \theta_3 = 1; \omega = I | \theta_1 = 1) = \frac{(1 - \pi)p(1 - p)^2}{\pi p + (1 - \pi)(1 - p)}$$

- Thus,  $v_i(1) = c$  is optimal for juror 1 if

$$2 \frac{\pi p^2(1 - p)}{\pi p + (1 - \pi)(1 - p)} \geq 2 \frac{(1 - \pi)p(1 - p)^2}{\pi p + (1 - \pi)(1 - p)}$$

$$\Rightarrow 2\pi p^2(1 - p) \geq (1 - \pi)p(1 - p)^2 + \pi p^2(1 - p)$$

# Sincere Voting Strategy

- ▶ After simplifying and rearranging, this inequality becomes

$$P_r(G | \theta_i=1, \theta_j=1, \theta_k=0) \stackrel{\text{(I am pivotal!)}}{=} \frac{\pi p^2(1-p)}{\pi p^2(1-p) + (1-\pi)p(1-p)^2} \geq \frac{1}{2}.$$

- ▶ LHS is just the **conditional** prob. **of guilt** given two signals of  $\theta = 1$  and one signal of  $\theta = 0$ .
- ▶ In other words, agent 1 wants to **vote to convict** if she believes that the defendant is more likely to be guilty than innocent, conditional on her signal and the belief that she is pivotal.
- ▶ Similarly, the requirement for a **vote of innocence** conditional on a signal of 0 is

$$\frac{\pi p(1-p)^2}{\pi p(1-p)^2 + (1-\pi)p^2(1-p)} \leq \frac{1}{2}.$$

- ▶ To sum, in any BNE in which voting corresponds to the private signals,
  1. Conditional on the supposition that i is pivotal and observes  $\theta_i = 1$ , the **posterior prob. of guilt** is greater than **1/2**; and
  2. Conditional on the supposition that i is pivotal and observes  $\theta_i = 0$ , the **posterior prob. of guilt** is less than **1/2**.

# Asymmetric Signal

- ▶ Thus, if sincere voting is incentive compatible, then

$$\frac{1-p}{p} \leq \frac{\pi}{1-\pi} \leq \frac{p}{1-p}.$$

- ▶ E.g., if  $\pi > p$ , then sincere voting is not incentive compatible.
- ▶ Under majority rule and symmetric signal precision (and equal prior  $\pi = 1/2$ ), sincere voting obtains in equilibrium (if  $p > 1/2$ ).
- ▶ Alternative way to obtain an *insincere/strategic voting equilibrium* is to introduce **asymmetric** signal:

$$\begin{aligned} p &\equiv \Pr(\theta_i = 1 | \omega = G), & q &\equiv \Pr(\theta_i = 0 | \omega = I), \\ 1-p &= \Pr(\theta_i = 0 | \omega = G), & 1-q &= \Pr(\theta_i = 1 | \omega = I), \end{aligned}$$

and we have here  $1 > p > q > 1/2$ .

- ▶ Then, the posterior probabilities (with equal prior  $\pi = 1/2$ ) are

$$\Pr[\omega = G | \theta_i = 1] = \frac{p}{p + (1-q)}, \quad \Pr[\omega = I | \theta_i = 0] = \frac{q}{(1-p) + q}.$$



# Strategic Voting Equ'm

- ▶ Define  $\sigma(s) \equiv$  prob. of voting one's signal,  $s = 0, 1$ .
- ▶ Typically, we have in equ'm;  $\sigma(1) \in (0, 1)$  and  $\sigma(0) = 1$ .
- ▶ Then,

$$\Pr[c|\omega = G] = p\sigma(1) + (1-p)(1-\sigma(0)) = p\sigma(1),$$

$$\Pr[a|\omega = G] = p(1-\sigma(1)) + (1-p)\sigma(0) = p(1-\sigma(1)) + (1-p),$$

$$\Pr[c|\omega = I] = (1-q)\sigma(1) + q(1-\sigma(0)) = (1-q)\sigma(1),$$

$$\Pr[a|\omega = I] = (1-q)(1-\sigma(1)) + q\sigma(0) = (1-q)(1-\sigma(1)) + q,$$

- ▶ Since the equ'm strategy requires randomization upon signal  $s = 1$ ,

$$\Pr[\omega = G|\theta_i = 1] \Pr[Piv|\omega = G] - \Pr[\omega = I|\theta_i = 1] \Pr[Piv|\omega = I] = 0,$$

where  $\Pr[Piv|\omega]$  is the prob. a vote is pivotal at state  $\omega$ :

$$\begin{aligned} \Pr[Piv|\omega = G] &= \binom{2}{1} \Pr[c|\omega = G] \Pr[a|\omega = G] \\ &= [p\sigma(1)][p(1-\sigma(1)) + (1-p)], \end{aligned}$$

$$\begin{aligned}\Pr[\text{Piv}|\omega = I] &= \binom{2}{1} \Pr[c|\omega = I] \Pr[a|\omega = I] \\ &= [(1 - q)\sigma(1)][(1 - q)(1 - \sigma(1)) + q]\end{aligned}$$

- ▶ Thus we solve for  $\sigma(1)$  from the above equation.
- ▶ Since  $\sigma(0) = 1$ , we finally check whether

$$\Pr[\omega = I|\theta_i = 0] \Pr[\text{Piv}|\omega = I] - \Pr[\omega = G|\theta_i = 0] \Pr[\text{Piv}|\omega = G] > 0$$

when  $\Pr[\text{Piv}|\omega]$  is evaluated at  $\sigma(1)$  that solves the indifference condition.

- ▶ For example, when  $p = 0.9$  and  $q = 0.6$ ,  $\sigma(1) = 0.9774$
- ▶ Under fixed  $(p, q)$ ,  $\sigma(1)$  typically decreases as  $n$  gets larger.

- ▶ Austen-Smith & Banks (1996) show that in many cases the sincere strategy is inconsistent with equilibrium behavior.
- ▶ It is easy to find parameters  $\pi$  and  $p$  for which one of the necessary conditions does not hold.
- ▶ There are alternative strategies jurors might choose.
- ▶ Jurors can randomize for some signals, vote the same way regardless of their signal, or use different strategies than other jurors use.
- ▶ Feddersen & Pesendorfer (1998) consider the properties of equilibria of this game when one varies the voting rule and number of jurors.

# Jury Voting with a Continuum of Signals

- ▶ Instead of receiving a binary signal, each juror now receives a signal  $\theta_i \in [0, 1]$  where  $\theta_i$  is drawn from a conditional distribution  $F(\theta_i|\omega)$ .
- ▶ This distribution function is associated with a different density function  $f(\theta_i|\omega)$  that satisfies the *monotone likelihood ratio* condition.
- ▶ A conditional density function satisfies the *strict monotone likelihood ratio condition* (SMLR) if  $\frac{f(\theta_i|G)}{f(\theta_i|I)}$  is a strictly monotone function of  $\theta_i$  on  $[0, 1]$ .
- ▶ To see why this assumption is important, note that Bayes' rule implies that

$$\begin{aligned}\Pr(G|\theta_i) &= \frac{f(\theta_i|G)\pi}{f(\theta_i|G)\pi + f(\theta_i|I)(1 - \pi)} \\ &= \frac{\frac{f(\theta_i|G)}{f(\theta_i|I)}\pi}{\frac{f(\theta_i|G)}{f(\theta_i|I)}\pi + (1 - \pi)}.\end{aligned}$$

- ▶ Accordingly,  $\Pr(G|\theta_i)$  is increasing in  $\theta_i$  if & only if  $f(\theta_i|G)/f(\theta_i|I)$  is increasing in  $\theta_i$ .
- ▶ Thus, the SMLR condition implies that higher signals correspond to higher posterior probabilities that  $\omega = G$ .

# Jury Voting with a Continuum of Signals

- ▶ To keep matters simple, we focus exclusively on symmetric strategies where voters who receive the same signal choose the same strategy.
- ▶ A symmetric strategy profile is, therefore, a mapping  $v_i(\theta_i) : [0, 1] \rightarrow \{c, a\}$ .
- ▶ As in the binary signal case, BNE strategies are those that are optimal when each agent acts conditionally on her private information and the conjecture that she is pivotal.
- ▶ An agent votes to convict if she thinks the prob. of guilt is no less than  $1/2$  and she votes to acquit if she thinks the prob. of guilt is no more than  $1/2$ .
- ▶ Because higher signals are better indicators of guilt, a natural conjecture is that the strategy must be weakly increasing.
- ▶ For low values of  $\theta_i$  an acquittal vote is cast and for high values of  $\theta_i$  a conviction vote is cast.

# Cut Point Strategy

- ▶ A monotone strategy of this form can be characterized by a cut point  $\hat{\theta} \in [0, 1]$ .
- ▶ Assume that agents  $i \in N \setminus i$  use the monotone strategy

$$v_i(\theta_i) = \begin{cases} c & \text{if } \theta_i \geq \hat{\theta} \\ a & \text{if } \theta_i < \hat{\theta} \end{cases}$$

- ▶ If all players other than  $i$  use this cut point strategy, the posterior prob. of  $\{\omega = G\}$  given signal  $\theta_i$  and the event that  $i$  is pivotal is given by

$$\begin{aligned} & \Pr(G|piv, \theta_i; \hat{\theta}) \\ &= \frac{\pi f(\theta_i|G)F(\hat{\theta}|G)^{N-r}[1 - F(\hat{\theta}|G)]^{r-1}}{\pi f(\theta_i|G)F(\hat{\theta}|G)^{N-r}[1 - F(\hat{\theta}|G)]^{r-1} + (1 - \pi)f(\theta_i|I)F(\hat{\theta}|I)^{N-r}[1 - F(\hat{\theta}|I)]^{r-1}} \end{aligned}$$

- ▶ This prob. is a function of the parameter  $\hat{\theta}$ .
- \* Here we assume  $r$ -rule, so we require  $r$  or more votes for conviction (majority rule if  $r = (N + 1)/2$  and unanimity rule if  $r = N$ ).

# Cut Point Equilibrium

- ▶ In this model the existence of a symmetric equ'm in which voters use a cut point hinges on finding a value of  $\hat{\theta}$  s.t.

$$\Pr(G|piv, \hat{\theta}; \hat{\theta}) = \frac{1}{2}$$

and demonstrating that  $\Pr(G|piv, \theta_i; \hat{\theta}) \leq \frac{1}{2}$  if  $\theta_i < \hat{\theta}$  and  $\Pr(G|piv, \theta_i; \hat{\theta}) \geq \frac{1}{2}$  if  $\theta_i > \hat{\theta}$ .

- ▶ Although analysis of examples is cumbersome, it is easy to derive conditions on the primitives of the game to ensure that such a  $\hat{\theta} \in (0, 1)$  exists.
- ▶ First,  $\Pr(G|piv, \theta_i; \hat{\theta}) \geq \frac{1}{2}$  if & only if

$$\begin{aligned} & \frac{\pi f(\theta_i|G) F(\hat{\theta}|G)^{N-r} [1 - F(\hat{\theta}|G)]^{r-1}}{(1 - \pi) f(\theta_i|I) F(\hat{\theta}|I)^{N-r} [1 - F(\hat{\theta}|I)]^{r-1}} \\ = & \frac{f(\theta_i|G)}{f(\theta_i|I)} \frac{\pi F(\hat{\theta}|G)^{N-r} [1 - F(\hat{\theta}|G)]^{r-1}}{(1 - \pi) F(\hat{\theta}|I)^{N-r} [1 - F(\hat{\theta}|I)]^{r-1}} \geq 1. \end{aligned}$$

# Existence of Cut Point Equilibrium

- ▶ SMLR then implies that if  $\Pr(G|piv, \hat{\theta}_i; \hat{\theta}) = 1/2$  then  $\theta_i < \hat{\theta}$  implies  $\Pr(G|piv, \theta_i; \hat{\theta}) \leq 1/2$  and  $\theta_i > \hat{\theta}$  implies  $\Pr(G|piv, \theta_i; \hat{\theta}) \geq 1/2$ .
- ▶ If  $\Pr(G|piv, 0; 0) \leq 1/2 \leq \Pr(G|piv, 1; 1)$  then the intermediate value theorem implies that such a cut point exists b/c the function  $\Pr(G|piv, \cdot; \cdot)$  is continuous.
- ▶ For a large class of games these boundary conditions are satisfied.
- ▶ In the simple binary signal model, equ'a where everyone uses the same rule and voting is determined by private information may not exist.
- ▶ This type of equ'm generally exists in the continuum model, however.
- ▶ Using the binary model, Feddersen & Pesendorfer (1998) show that the unanimity rule is a uniquely bad way to aggregate information for large populations b/c in equ'm voters condition on the assumption that everyone else is voting to convict.
- ▶ In the continuum model, Meirowitz (2002) shows that the unanimity rule often turns out to be as good as the other voting rules.



# Voluntary Voting Model

- ▶ Two candidates, A and B, in majority voting election.
- ▶ Two equally likely states of nature,  $\alpha$  and  $\beta$ .
- ▶ A is the better choice in state  $\alpha$  and B, in state  $\beta$ .
  - In state  $\alpha$ , payoff is 1 if A is elected and 0 if B is elected; vice versa in state  $\beta$ .
- ▶ The size of the electorate is a random variable, distributed according to a *Poisson* distribution with mean  $n$ .
  - The probability that there are exactly  $m$  voters is  $e^{-n} n^m / m!$ .
- ▶ Prior to voting, each voter receives a private signal  $S_i$  regarding the true state of nature, either  $a$  or  $b$ ;  $\Pr[a|\alpha] = r$  and  $\Pr[b|\beta] = s$ ; the posteriors given by

$$q(\alpha|a) = \frac{r}{r + (1 - s)}, \quad q(\beta|b) = \frac{s}{s + (1 - r)}.$$

- $r \geq s > 1/2$  implies  $q(\alpha|a) \leq q(\beta|b)$ .

# Pivotal Events

- ▶ Event  $(j, k)$ ,  $j$  votes for A and  $k$  votes for B.
- ▶ An event is *pivotal* for A if a single additional vote for A changes the outcome, written  $Piv_A$ .
- ▶ Under majority rule, one additional vote for A makes a difference only if (i) there is a tie; or (ii) A has one vote less than B.

$$T = \{(k, k) : k \geq 0\}, \quad T_{-1} = \{(k-1, k) : k \geq 1\}, \quad Piv_A = T \cup T_{-1}$$

- ▶ Similarly,  $Piv_B = T \cup T_{+1}$ ,  $T_{+1} = \{(k, k-1) : k \geq 1\}$ .
- ▶  $\sigma_A, \sigma_B$  are the *expected* number of votes for A, B in state  $\alpha$ ;  $\tau_A, \tau_B$  are the *expected* number of votes for A, B in state  $\beta$ .
- ▶ With abstention allowed,  $\sigma_A + \sigma_B \leq n$ ,  $\tau_A + \tau_B \leq n$  (equality w/o abstention).

# Pivotal Events

- ▶ If the realized electorate is of size  $m$  with  $k$  votes for A and  $l$  votes for B ( $m - k - l$  abstention),

$$\Pr[(k, l)|\alpha] = e^{-\sigma_A} \frac{\sigma_A^k}{k!} e^{-\sigma_B} \frac{\sigma_B^l}{l!}.$$

- \* For the probability of the event  $(k, l)$  in state  $\beta$ , replace  $\sigma$  by  $\tau$ .

$$\Pr[T|\alpha] = e^{-\sigma_A - \sigma_B} \sum_{k=0}^{\infty} \frac{\sigma_A^k}{k!} \frac{\sigma_B^k}{k!},$$

$$\Pr[T_{-1}|\alpha] = e^{-\sigma_A - \sigma_B} \sum_{k=1}^{\infty} \frac{\sigma_A^{k-1}}{(k-1)!} \frac{\sigma_B^k}{k!},$$

$$\Pr[\text{Piv}_A|\alpha] = \frac{1}{2} \Pr[T|\alpha] + \frac{1}{2} \Pr[T_{-1}|\alpha]$$

where  $\text{Piv}_A = T \cup T_{-1}$  is the set of events where one additional vote for A is decisive, and we have the coefficient  $1/2$  because the additional vote for A breaks a tie or leads to a tie.

# Pivotal Events

- ▶ Similarly,

$$\Pr[\text{Piv}_B|\beta] = \frac{1}{2} \Pr[T|\beta] + \frac{1}{2} \Pr[T_{+1}|\beta]$$

where  $\text{Piv}_B = T \cup T_{+1}$  is the set of events where one additional vote for B is decisive.

- ▶ Define *modified Bessel functions*

$$I_0(z) = \sum_{k=0}^{\infty} \frac{(z/2)^k}{k!} \frac{(z/2)^k}{k!}, \quad I_1(z) = \sum_{k=1}^{\infty} \frac{(z/2)^{k-1}}{(k-1)!} \frac{(z/2)^k}{k!}$$

and rewrite the probabilities of close elections in terms of these functions

$$\begin{aligned} \Pr[T|\alpha] &= e^{-\sigma_A - \sigma_B} I_0(2\sqrt{\sigma_A \sigma_B}) \\ \Pr[T_{\pm 1}|\alpha] &= e^{-\sigma_A - \sigma_B} \left(\frac{\sigma_A}{\sigma_B}\right)^{\pm 1/2} I_1(2\sqrt{\sigma_A \sigma_B}). \end{aligned}$$

- ▶ For  $z$  large, we also have

$$I_0(z) \approx \frac{e^z}{\sqrt{2\pi z}} \approx I_1(z).$$

# Compulsory Voting

- ▶ By compulsory voting each voter must cast a vote for either A or B.
- ▶ Vote sincerely in compulsory voting equilibrium?
- ▶ Given sincere and compulsory voting,  $\sigma_A = nr$ ,  $\sigma_B = n(1 - r)$ ,  $\tau_A = n(1 - s)$ ,  $\tau_B = ns$ .
- ▶ As  $n$  increases, both  $\sigma \rightarrow \infty$ ,  $\tau \rightarrow \infty$ , and so the previous approximations for  $I_0(z)$ ,  $I_1(z)$  imply

$$\frac{\Pr[\text{Piv}_A|\alpha] + \Pr[\text{Piv}_B|\alpha]}{\Pr[\text{Piv}_A|\beta] + \Pr[\text{Piv}_B|\beta]} \approx \frac{e^{2n\sqrt{r(1-r)}}}{e^{2n\sqrt{s(1-s)}}} \times K(r, s)$$

where  $K(r, s)$  is positive and independent of  $n$ .

- ▶  $r > s > 1/2$  also implies  $s(1 - s) > r(1 - r)$  and so RHS goes to zero as  $n$  increases.

# Compulsory Voting

- ▶ This implies that, when  $n$  is large and a voter is pivotal, state  $\beta$  is infinitely more likely than state  $\alpha$ .
- ▶ Thus, voters with  $a$  signals will not wish to vote sincerely.

**Proposition 1:** *Suppose  $r > s$ . If voting is compulsory, sincere voting is not an equilibrium in large elections.*

- ▶ This result also holds for a fixed number of voters (Feddersen & Pesendorfer APSR 1998).

# Voluntary Voting

- ▶ Costly voting: one's cost of voting is private info and an independent draw from a continuous distribution  $F$  with support  $[0, 1]$  -  $F$  admits a density  $f > 0$  on  $[0, 1]$ .
- ▶ Voting costs are independent of the signals.
- ▶ There exists an equilibrium of this voluntary (and costly) voting game with the following features;
  - (i) There exists a pair of positive *threshold costs*  $c_a, c_b$  s.t. a voter with cost  $c$  and signal  $i = a, b$  votes (does not abstain) if & only if  $c \leq c_i$ . The threshold costs determine differential *participation rates*  $F(c_a) = p_a, F(c_b) = p_b$ .
  - (ii) All those who vote do so sincerely - i.e., all those with signal  $a$  vote for A and those with signal  $b$  vote for B.

# Equ'm Participation Rates

- ▶ We show that when all those who vote do so sincerely, there is an equ'm in cutoff strategies.
- ▶ There exists a threshold cost  $c_a > 0$  ( $c_b > 0$ ) s.t. all voters with signal  $i$  and cost  $c \leq c_a$  ( $c \leq c_b$ ) go to the polls and vote for A (B).
- ▶ These then determine participation probabilities  $p_a = F(c_a)$ ,  $p_b = F(c_b)$  for voters with signal  $a$ ,  $b$ , respectively.
- ▶ Now the expected numbers of votes for A, B in state  $\alpha$  are  $\sigma_A = nrp_a$ ,  $\sigma_B = n(1-r)p_b$ ; and those in state  $\beta$  are  $\tau_A = n(1-s)p_a$ ,  $\tau_B = nsp_b$ , respectively.
- ▶ We look for participation rates  $p_a$ ,  $p_b$  s.t. a voter with signal  $a$  and cost  $c_a = F^{-1}(p_a)$  is indifferent b/w going to the polls and staying home;

$$(IRa) \quad U_a(p_a, p_b) \equiv q(\alpha|a) \Pr[\text{Piv}_A|\alpha] - q(\beta|a) \Pr[\text{Piv}_A|\beta] = F^{-1}(p_a)$$



# Equ'm Participation Rates

where the pivot probabilities are determined using the expected vote totals  $\sigma, \tau$ .

- ▶ Similarly, a voter with signal  $b$  and cost  $c_b = F^{-1}(p_b)$  must also be indifferent;

$$(IRb) \quad U_b(p_a, p_b) \equiv q(\beta|b) \Pr[Piv_B|\beta] - q(\alpha|b) \Pr[Piv_B|\alpha] = F^{-1}(p_b).$$

**Proposition 2:** *There exist participation rates  $p_a^* \in (0, 1)$  and  $p_b^* \in (0, 1)$  that simultaneously satisfy (IRa) and (IRb).*

- ▶ Intuition for positive participation rates: assume  $p_a = 0$ .
- ▶ Then the only pivotal events are  $(0, 0)$  and  $(0, 1)$ .

# Equ'm Participation Rates

- ▶ Hence conditional on being pivotal

$$\frac{\Pr[\text{Piv}_A|\alpha]}{\Pr[\text{Piv}_A|\beta]} = \frac{e^{-n(1-r)p_b}}{e^{-nsp_b}} \times \frac{1 + n(1-r)p_b}{1 + nsp_b}.$$

- ▶ The ratio of the exponential terms favors state  $\alpha$  while the ratio of the linear terms favors state  $\beta$ ; and the exponential terms always dominate.
- ▶ Since state  $\alpha$  is perceived more likely than  $\beta$  by a voter with signal  $a$  who is pivotal, the payoff from voting is positive.
- ▶ We also have

**Lemma 1:** *If  $r > s$ , then any solution to (IRa) and (IRb) satisfies  $p_a^* < p_b^*$ , with equality if  $r = s$ .*

# Sincere Voting

- ▶ Given the (equ'm) participation rates, we can show that it is a best-response for every voter to vote sincerely.
- ▶ We begin with a lemma;

**Lemma 2:** *If voting behavior is s.t.  $\sigma_A > \tau_A$  and  $\sigma_B < \tau_B$ , then*

$$\frac{\Pr[\text{Piv}_B|\alpha]}{\Pr[\text{Piv}_B|\beta]} > \frac{\Pr[\text{Piv}_A|\alpha]}{\Pr[\text{Piv}_A|\beta]}.$$

- ▶ On the set of “marginal” events where the vote totals are close (i.e., a voter is pivotal), A is more likely to be leading in state  $\alpha$  and more likely to be trailing in state  $\beta$ .
- ▶ Let  $(p_a^*, p_b^*)$  be equ'm participation rates.
- ▶ A voter with signal  $a$  and cost  $c_a^* = F^{-1}(p_a^*)$  is just indifferent b/w voting and staying home;

$$(IRa) \quad q(\alpha|a) \Pr[\text{Piv}_A|\alpha] - q(\beta|a) \Pr[\text{Piv}_A|\beta] = F^{-1}(p_a^*).$$

# Sincere Voting

- ▶ To show: sincere voting is optimal for a voter with signal  $a$  if others are voting sincerely;

$$\begin{aligned} (ICa) \quad & q(\alpha|a) \Pr[Piv_A|\alpha] - q(\beta|a) \Pr[Piv_A|\beta] \\ & \geq q(\beta|a) \Pr[Piv_B|\beta] - q(\alpha|a) \Pr[Piv_B|\alpha]. \end{aligned}$$

- ▶ LHS is the payoff from voting for A whereas RHS is the payoff to voting for B.
- ▶  $p_a^* > 0$  combined with (IRa) implies

$$\frac{\Pr[Piv_A|\alpha]}{\Pr[Piv_A|\beta]} > \frac{q(\beta|a)}{q(\alpha|a)}.$$

- ▶ Then by Lemma 2,

$$\frac{\Pr[Piv_B|\alpha]}{\Pr[Piv_B|\beta]} > \frac{q(\beta|a)}{q(\alpha|a)}.$$

- ▶ But then, the last inequality is equivalent to

$$q(\beta|a) \Pr[Piv_B|\beta] - q(\alpha|a) \Pr[Piv_B|\alpha] < 0.$$

- ▶ Similarly, we combine  $p_b^* > 0$ , Lemma 2, and

$$(IRb) \quad q(\beta|b) \Pr[Piv_B|\beta] - q(\alpha|b) \Pr[Piv_B|\alpha] = F^{-1}(p_b^*)$$

to show

$$(ICb) \quad \begin{aligned} & q(\beta|b) \Pr[Piv_B|\beta] - q(\alpha|b) \Pr[Piv_B|\alpha] \\ & \geq q(\alpha|b) \Pr[Piv_A|\alpha] - q(\beta|b) \Pr[Piv_A|\beta]. \end{aligned}$$

**Proposition 3:** *Under voluntary participation, sincere voting is incentive compatible.*

- ▶ We can also show that all equ'a involve sincere voting (Krishna & Morgan JET 2012).